

# Choosing Stress Scenarios for Systemic Risk Through Dimension Reduction

Matthew Pritsker\*

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## Abstract

Regulatory stress-testing is an important tool for ensuring banking system health in many countries around the world. Current methodologies ensure banks are well capitalized against the scenarios in the test, but it is unclear how resilient banks will be to other plausible scenarios. This paper proposes a new methodology for choosing scenarios that uses a measure of systemic risk with Correlation Pursuit variable selection, and Sliced Inverse Regression factor analysis, to select variables and create factors based on their ability to explain variation in the systemic risk measure. The main result is under appropriate regularity conditions, when the banking system is well capitalized against stress-scenarios based on movements in the factors, then an approximation of systemic risk is low, i.e. the banking system will be well capitalized against the other plausible scenarios that could affect it with high probability. The paper also shows there are circumstances when several scenarios may be required to achieve systemic risk objectives. The methodology should be useful for regulatory stress-testing of banks. Although not done in this paper, the methodology can potentially be adapted for stress-testing of other financial firms including insurance companies and central counterparties.

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\*The authors is a member of the Risk and Policy Analysis Unit at the Federal Reserve Bank of Boston. Matt Pritsker's contact information is as follows: ph: (617) 973-3191, email: matthew.pritsker at bos.frb.org. The views expressed in this paper are those of the author but not necessarily those of the Federal Reserve Bank of Boston or other parts of the Federal Reserve System. The author thanks Maggie Li for excellent research assistance, and John Ammer, Yasushi Hamao, Christian Julliard, Chen Zhou, Eric Schaanning, and Rama Cont for valuable comments. The author also thanks seminar participants at the Federal Reserve Bank of New York, the Federal Reserve Bank of Chicago, the Federal Reserve Bank of Boston, the Federal Reserve Board, the International Monetary Fund, the Bank of England, and the University of Sussex, as well as conference participants at the London School of Economics, the Midwest Finance Association, the Conference on Computational and Financial Econometrics, the Paris Financial Management Association Conference, the Conference on Financial Stability and Macroprudential Policy at the Norges Bank, and at the Econometric Society European Meetings. The author also thanks Jens Christensen and Jonathan Wright for invaluable help on term-structure modeling aspects of this work. All errors are the responsibility of the author.

# 1 Introduction

The great recession that accompanied the Financial Crisis of 2007-2009 underlined the role of the financial sector in real economic activity, and it highlighted the importance of controlling systemic risk, the risk that many banks and financial institutions become financially distressed at the same time and thus become impaired in their ability to provide financial intermediation for the real sector. As part of the US regulatory response to the financial crisis, stress-tests were conducted that assessed the ability of individual banks to maintain sufficient capital (measured by net worth over risk-weighted assets) to perform as financial intermediaries during a small number of adverse hypothetical macroeconomic and financial scenarios. Remedial action was required for banks that had capital shortfalls. Since the first U.S. stress test in 2009, regulatory stress testing has become the primary tool to assess the capital adequacy of many banks in the US. In addition, it is used as a regulatory tool in many other countries.

Regulatory stress scenarios are designed to cover the risks in banks loan books. For banks with significant trading operations, separate stress scenarios are often used to cover the risks in their trading books.<sup>1</sup> An important goal of regulatory stress-testing is to ensure that systemic risk is low. This requires the financial system to be well capitalized against the small number of scenarios in the stress-test *and* against the much broader set of likely scenarios that the financial system may face. To accomplish this objective using a small number of stress scenarios it is necessary to choose the scenarios well. This involves selecting the right set of variables to stress, and the appropriate directions and magnitudes in which to stress them. Failure to satisfy these necessary conditions can result in a missing variables problem if a regulatory scenario stresses variables banks are not exposed to while failing to stress important variables that they are exposed to, in a missing directions problem if the stress scenario is in a direction in which banks are hedged or make profits while failing to stress directions in which banks make losses, or in an inadequate magnitude problem if the variables that banks have significant exposures to are stressed by an insufficient amount.

As an example of the missing variable problem, regulatory scenarios for banks' loan books are often formulated on the basis of a small number of macroeconomic and financial variables that only weakly explain banks' P&L [Guerrieri and Welch (2012), Bolotnyy et. al (2015)] and whose movements often fail to forecast financial crises since such crises often occur before the macroeconomy turns down [(Borio et al (2012), Alfaro and Drehmann (2009)].<sup>2</sup> These findings suggest that additional variables may be needed to capture the risk in banks' loan books. As an example of the missing directions problem, capital adequacy for large banks' trading book positions is often assessed on the basis of a few regulatory scenarios that specify the movements of many

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<sup>1</sup>For details see Federal Reserve Board (2014), European Banking Authority (2014), and Bank of England (2014).

<sup>2</sup>The Federal Reserve's 2015 CCAR stress scenario for the 6 largest banks trading book positions stresses tens of thousand of variables. But, only 16 macro-financial variables were utilized for U.S. banking book exposures, and only 12 macro-financial variables were used for banking book exposures outside of the United States [Federal Reserve Board (2014)].

variables for EBA/UK regulatory stress tests, and for a very large number of variables (10,000+) for US regulatory stress tests. In such a high dimensional setting, unless the scenario is chosen very carefully, important directions of risk-taking may be missed. Moreover, in such a high dimensional setting, it is not apriori straightforward how to choose the correct magnitudes by which to stress the variables.

To overcome the missing variable and missing direction problems, U.S. regulations also require each bank to construct a scenario that stresses its most important vulnerabilities. However, because this approach is based on bank-specific vulnerabilities, it does not ensure banks are well capitalized against common vulnerabilities, and hence cannot ensure that systemic risk, the risk that many banks in the financial system becomes undercapitalized together, is sufficiently low.

This paper introduces a new methodology for creating regulatory stress scenarios; it chooses the variables to use in a stress-scenario, and the directions and amounts the variables need to be moved so that if banks are well capitalized against the scenario, then under some conditions (discussed below) systemic risk will be low as measured by an approximate systemic risk objective. The new methodology relies on dimension reduction techniques. It is premised on the idea although the value of banks positions are driven by many variables, these variables are driven by a smaller number of latent economic factors. The factors are assumed to be the most important determinant of banks risk at a portfolio level. At an individual bank level, its risk is not determined by all of the factors, but only those it has not hedged against. On an economy wide level, banks' common exposures to unhedged factors can cause them to experience joint distress and are hence a source of systemic risk. This reasoning suggests that for stress tests to keep systemic risk low, stress testing policy needs to ensure banks remain well capitalized against movements in the most important factors that explain their joint distress. The methodology in this paper pursues this idea by illustrating how to use supervisory information and statistical techniques to identify banks exposures to the unhedged factors, and to then design stress scenarios that keep systemic risk low.

To identify the factors, the variables that affect banks portfolios are simulated, then using supervisory information on risk exposures, changes in the variables are mapped into changes in the value of banks' portfolios, and into a measure of banks' joint financial distress. Then a principal components factor analysis using Sliced Inverse Regression (SIR) [Li, (1991)], is conducted that identifies an orthogonalized set of risk factors based on their ability to explain banks' joint distress.<sup>3</sup> Because banks' joint distress in the simulations can only depend on economic factors that banks have not fully hedged against, the factors identified by SIR will only depend on the unhedged factors.<sup>4</sup> As explained below, SIR is more accurate if it does not rely on too many variables to create the factors. To choose the variables that should be used for SIR, it is assumed that the

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<sup>3</sup>Technically, the factors are chosen based on their correlation with an optimally chosen transformation of banks joint distress.

<sup>4</sup>Formally, the factors identified by SIR will be spanned by banks unhedged factors, and in some conditions both sets of factors will span the same space.

information on banks joint distress that is contained in all of the variables, is also contained in a smaller subset of variables that are best for creating factors that explain banks' joint distress. The best variables for creating the factors in a SIR framework are chosen using Correlation Pursuit (COP) [Zhong et al. (2012)] variable selection.

To construct a stress scenario using the factors, the factors are shocked by chosen amounts, and then all variables are set to their conditional expected values given the factor shocks. By relying on systemic risk factors to determine how variables are shocked, the scenario by construction moves variables in stressful directions from a systemic risk perspective. The shock sizes and directions in the stress-scenario are chosen so that if banks are well capitalized against the scenario, then (if feasible) regulators systemic risk objective will be achieved. Because the approach in this paper chooses one or a small number of scenarios in order to satisfy a systemic risk objective, I refer to the approach in the paper as the Systemically Chosen Scenario Approach, or SCSA.

The SIR and closely related COP method are both based on factor analysis, but differ from classical factor analysis. The purpose of classical factor analysis is to summarize the information about the joint behavior of a large number of variables by a much smaller number of factors that are linear combinations of the variables. By contrast, SIR and COP are supervised factor analysis methods that creates factor based on their ability to explain a dependent variable. Partial Least Squares (PLS) is another method for performing supervised factor analysis. SIR and COP differ from PLS because the latter typically requires the left hand side variable to be a linear combination of the factors, while COP and SIR allow the dependent variable to be a nonlinear function of the factors. Because joint distress caused by asset bubbles bursting, or asset fire sales may display nonlinear dynamics, an advantage of using SIR and COP to identify the factors is that these methodologies may still be able to uncover the factor structure even when there are non-linearities.<sup>5</sup>

This paper contributes to both the practice and theory of regulatory stress-testing.<sup>6</sup> Current regulatory practice uses different approaches to specify scenarios for banks' loan and trading books. Loan book stresses are usually based on macroeconomic models, and consistent with those models, utilize a relatively small number of macroeconomic and financial variables. As noted earlier, these variables only weakly explain banks P&L, suggesting additional variables may be needed.<sup>7</sup> By contrast, in the trading book, regulators specify stress scenarios using a very large number of variables based on the variables banks use to model their risks. This approach to trading book scenarios

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<sup>5</sup>Gibson and Pritsker (2001), and Giglio et al (2012) use PLS in the context of dimension reduction for risk measurement. A novel aspect of Giglio et al is it uses quantile regression with partial least squares, and hence is an early attempt to estimate a nonlinear factor structure using PLS. In addition, Kelly and Pruitt (2013) use PLS for stock market prediction, and Groen and Kapetanios discuss its use for macroeconomic forecasting.

<sup>6</sup>Seminal contributions to systemic risk stress-testing include the macro-financial models of the Bank of Austria [Boss et. al (2006)] and the Bank of England [Alessandri et. al (2009)]. Bookstaber et. al. (2013) and Schuermann (2013) provide critical reviews of the literature and regulatory practice.

<sup>7</sup>Recent research suggests that the clustering of defaults is not explained by macro and fincaill variables, that the variables which explain this clustering have not yet been identified [Das et al (2007); Azizpour et al (2015)].

mitigates the missing variables problem, but exacerbates difficulties in choosing the direction and magnitude by which trading book variables should be shocked in a stress scenario. The methods advocated in this paper have potential to improve stress testing for both sets of books. To improve loan book variable selection this paper advocates expanding the set of variables used to model loan book risks to include the full set of variables banks use to model those risks, and then choose the loan book variables that are most useful for modeling systemic risk. To address the missing directions problem in both the loan and trading books, the paper uses dimension reduction to identify the direction and magnitude of banks vulnerabilities to a smaller number of identified systemic risk factors, and chooses stress scenarios to achieve regulatory objectives based on this information.

This paper is related to a growing literature on systemic risk measurement [Bisias et. al.(2012)]. The methodology in this paper does not address all types of systemic risk, but it is related to banks becoming financially distressed by becoming under-capitalized together. The approach to derive stress scenarios in this paper requires a measure of systemic risk based on banks' joint distress. Measures that could be used in this paper include aggregated across banks versions of Systemic Expected Shortfall (SES) [Acharya et al, 2010], the Distressed Insurance Premium (DIP) [Huang et al (2009)], or System Assets in Distress (SAD) [Pritsker, (2014)].<sup>8</sup> Although not pursued in this paper, it should be possible to alternatively measure systemic risk based on commonality in financial institutions' liquidity mismatches, and then identify factors that explain vulnerability to this commonality.<sup>9</sup>

This paper is closely related to a few papers in the stress-testing literature. Pritsker (2014) proposes a methodology to achieve systemic risk objectives at lowest capital cost in a framework that uses a very large number of stress scenarios. This paper proposes a complementary approach that attempts to accomplish a similar objective by utilizing a smaller number of stress scenarios that are very carefully chosen. Reliance on a few scenarios is more consistent with regulatory practice, and may be more practical to implement if computing many scenarios is too costly. Another complementary paper is Kapinos and Mitnik (2014). They use LASSO regression and factor analysis to improve on variable selection and P & L modeling as part of stress-testing. Their application differs from this paper in several important ways, most importantly they do not choose variables or factors based on their ability to explain systemic risk, and they do not choose stress scenarios to achieve a systemic risk objective.<sup>10</sup>

An important contribution of this paper is that stress-scenarios are chosen to achieve a particular

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<sup>8</sup>System Assets in Distress is aggregated across banks.

<sup>9</sup>See BIS 2013 for a discussion of liquidity stress testing.

<sup>10</sup>Kapinos and Mitnik (KM) (2014) choose which macro-variables and transforms of macro-variables explain components of banking-book P&L using LASSO regression; they then apply principal components to the chosen variables and use the components to model how banks respond to changes in macro variables as part of stress-testing. Although this paper and KM are similar in choosing variables and factors, their focuses are different. KM chooses variables to estimate models that relate the variables to banks P&L. By contrast, this paper takes the relationships between the variables and banks P&L as given; it then uses these relationships to choose variables and factors to create stress-scenarios to attain systemic risk objectives.

systemic risk objective. Some papers on stress scenario selection pursue a related approach for firm stress-testing by choosing a stress scenario that generates the largest losses for the firm from among a set of possible scenarios. If the firm is well capitalized against the worst-case scenario, it is well capitalized against all scenarios in the set with a probability exceeding the probability of the set [Breuer et al (2009), Flood and Korenko (2013)].<sup>11</sup> An advantage of this worst-case approach is if the objective for a firm is to achieve capital adequacy with a given probability, then if the set has that probability, assuring capital adequacy against the set meets and exceeds the objective. A disadvantage of worst case approaches is the objective is often exceeded by large amounts requiring a firm to hold far more capital than is needed. The method in this paper achieves the objective but is less conservative because it is designed to just satisfy the objective, not exceed it. An additional contribution of the approach in this paper is the objective is not for one firm, but is instead based on a systemic risk objective function for the economy.

The rest of the paper proceeds in five sections. Section 2 illustrates potential pitfalls in regulatory stress testing that the SCSA approach is designed to address. Section 3 explains the SCSA methodology. Section 4 illustrates the SCSA methodology for some stylized trading portfolios that are exposed to interest-rate and stock market risk. Section 5 discusses extensions to account for exposure uncertainty and for uncertainty about the shape of interbank networks. A final section concludes.

## 2 Pitfalls in regulatory stress testing.

This section uses simple examples to illustrate how if regulatory scenarios are not chosen well, it may be difficult to achieve systemic risk objectives. The missing variables problem is straightforward. Therefore, the exposition below focuses on missing directions, and the requirement that banks must also create their own bank specific scenarios.

### Missing Stress-Test Directions

This section illustrates that if stress-scenarios are created without accounting for the directions in which banks take risk, the stress-testing exercise may fail to require banks to hold capital against potentially very significant risk exposures. If, in addition, the direction of missed risk is common across banks, the fact that the stress-scenario missed the risk-taking can itself be a source of systemic risk. These points are simple to illustrate in a univariate setting. For example, suppose all banks write call options on the S&P 500 stock index, and that is the only asset position that they have. In this setting banks are only subject to the risk that the stock market moves in an upward direction. Because all banks have a common exposure, the chance of large upward movements represents a

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<sup>11</sup>An alternative approach chooses the most likely scenario that generates losses of a given amount [Glasserman, Kang, and Kang (2015)].

systemic risk. If there is only one regulatory scenario and it posits that stock prices move down, it would have chosen the wrong stress direction, and not required banks to hold additional capital despite the systemic risk of their positions.

The example of missing directions when there is only one asset is contrived and unrealistic. But, analogous problems exist in higher dimensional settings. A slightly more complicated example involves a setting in which banks are only exposed to two assets. In this setting, suppose the stress scenario takes the form that asset 1 goes down 5% and asset 2 goes down 10%, and suppose for simplicity that banks perfectly anticipate that this will be the scenario.<sup>12</sup> A bank can perfectly hedge this scenario by going long \$2 million in asset 1 and short \$ 1 million in asset 2. Such a portfolio will lose nothing in the stress scenario, even though the portfolio has risk.<sup>13</sup> If the scale of the portfolio is increased 1000-fold to long \$2 billion in asset 1 and short \$ 1 billion in asset 2, it still loses nothing in the scenario even though the second portfolio is much riskier than the first. This two asset example illustrates that in higher dimensional settings, it is very simple to create portfolios that lose nothing in the stress-scenario provided the scenario is anticipated.<sup>14</sup> Even if a scenario is unanticipated, the stress-test may still miss important directions of banks risk-taking, especially in banks' trading book where the dimension of risk-taking is potentially very high. Similar issues also arise in banks' loan books, but the extent of the problem depends on the dimensionality and type of risk-taking in banks loan books.<sup>15</sup>

## Bank Specific Stress Tests

As noted in the introduction, a possible method to address the missing variable and stress direction problems is to allow each bank to conduct its own bank-specific stress-test that is tailored to the exposures that it deems to be most important. An inherent difficulty with this approach is by focusing on bank specific risks, it may fail to ensure the banking system is well enough capitalized against systemic risk. To illustrate this idea, suppose there are  $M$  banks that are stress-tested and that failure or collective financial weakness at  $N$  of these banks could pose systemic risk problems. Further suppose that  $N$  banks are exposed to a common factor  $f$  and idiosyncratic

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<sup>12</sup>Glasserman and Tangirala (2015) provide evidence that the CCAR stress scenarios have been predictable from year to year.

<sup>13</sup>It loses money if asset 1 goes down in value and asset 2 goes up in value.

<sup>14</sup>In higher dimensional settings missing direction problems are even worse. For example, if there are  $N$  assets, and their returns in the stress scenario are  $r_1, r_2, r_3, \dots, r_N$  then there are  $N - 1$  linearly independent portfolios that lose nothing in the scenario: a portfolio that is long  $r_2$  dollars of asset 1, and short  $r_1$  dollars of asset 2 loses nothing; a portfolio that is long  $r_3$  dollars of asset 2 and short  $r_2$  dollars of asset 3 loses nothing; ... a portfolio that is long  $r_N$  dollars of asset  $N - 1$  and short  $r_{N-1}$  dollars of asset  $N$  lose nothing. Additionally, all linear combinations of these  $N - 1$  portfolios lose nothing in the stress scenario.

<sup>15</sup>For example, if risks in the loan book only depend on a single macroeconomic factor, then it is only important to choose whether to move this variable up or down. If instead loan book risks are driven by multiple factors, then it is again important to choose the stress directions well. As one simple example, if a substantial part of banks domestic loans were denominated in foreign currencies, as has been the case in some parts of Europe [Yesin(2013)], then the borrowers may default if the domestic currency depreciates against the currencies in which the loans are made. In this circumstance, capturing exchange rate risk, and moving exchange rates in the correct directions in the stress scenario is important. Similarly, if banks have a significant amount of loans to energy industry firms, then choosing the direction in which oil prices and the macro-factors move is important.

risks  $\epsilon_n$ ,  $n \in 1, \dots, N$  that are independent across banks. From a bank specific perspective, if their idiosyncratic risk is more volatile than their  $f$  risk, then bank specific tests would be tailored around each banks idiosyncratic risk, and banks would choose to hold capital against their idiosyncratic risk. However, from a systemic perspective it is the  $f$  risk that is more important because it causes more banks to experience trouble together. Hence it may merit banks holding more capital against the  $f$  risk, and hence a relatively large movement in  $f$  for stress-testing purposes. But, if  $f$  is an omitted factor in regulatory scenarios, and is not chosen in bank-specific scenarios because it is relatively unimportant for individual banks, then the bank specific stress-tests will not ensure banks are capitalized against  $f$  and will hence potentially miss an important systemic risk factor.

In sum, this section has illustrated why it is important for regulatory stress-scenarios to specify the correct variables and correct stress-directions, and it has illustrated that bank specific scenarios are not necessarily a complete solution for missing variables problems in regulatory stress-testing.

### 3 The SCSA methodology

As noted above, current regulatory stress-scenario selection has four main shortcomings:

1. Banks exposures are not formally used in scenario selection.
2. The wrong variables may be utilized in scenario formulation.
3. The variables may be stressed in the wrong directions.
4. The stress-scenarios are not explicitly designed to achieve a systemic risk objective.

The SCSA methodology helps to address all four problems. It is based on three principles:

**Principle 1** *The value of banks positions (assets, liabilities, and derivative securities) depends on a large number of variables including interest rates, FX rates, stock returns, implied volatilities, etc. The large number of variables in turn depend on a much smaller number of underlying potentially latent economic factors.*

**Principle 2** *Systemic impairment is the event that too many banks become financially distressed during the same period of time. Systemic risk is the probability that systemic impairment occurs. One way systemic impairment can occur is if banks are exposed to common economic factors and those factors move in ways that are unfavorable to many banks at the same time.*

**Principle 3** *Regulatory stress scenarios should be chosen so that if the banking system is well capitalized against the stress-scenarios, then systemic risk is low with high probability.*



Principles 1 and 2 suggest that systemic impairment can occur if the common factors that banks are exposed to move against them by enough to cause the banks to jointly experience financial distress. This suggests, by principle 3, that stress-scenarios should be designed based on movements in the factors.

The remaining analysis is divided in two parts: section 3.1, provides economic and statistical theory to identify the factors; section 3.2 provides details on how to use the identified factors to choose stress scenarios, and if feasible to ensure systemic risk is low.

### 3.1 Identifying Systemic-Risk Factors using SCSA

The goal of this section is to illustrate how to estimate common factors that affect the value of banks portfolios based on the set of tangible variables that banks included in a stress test use to model their assets, plus additional tangible variables that regulators may use to use value banks assets.<sup>16</sup> The total set of variables is denoted by the  $1 \times N$  vector  $X$ .

By principle 1, it is assumed that  $X$  is driven by a factor structure,

#### Assumption 1

$$X = G(F_A, F_B, U) \tag{1}$$

where  $G(\cdot)$  is a function linking variables to the factors, and  $F_A$  and  $F_B$  are  $1 \times K_A$  and  $1 \times K_B$  vectors of potentially latent economic factors, and  $U$  is a  $1 \times N$  vector of idiosyncratic risks that are independent of the factors.<sup>17</sup>

In this setting, the factors  $F_B$  denote factors that all banks are not exposed to or fully hedge against, while  $F_A$  represent factors that banks remain exposed to.<sup>18</sup> To avoid difficulties with missing variables, it is important not to create a stress-test based on  $F_B$ : since all banks are hedged against  $F_B$  nothing would be learned about banks systemic risk by stressing  $F_B$ .<sup>19</sup> Conversely,  $F_A$  represents common factors that banks remain exposed to; by principle 2 stress-tests should be based on those factors. Each bank  $i$ 's remaining idiosyncratic risk after hedging is represented by

<sup>16</sup>Intangible variables include a loan officers judgment.

<sup>17</sup>The space that the factors span can be identified without knowledge of  $G(\cdot)$ .

<sup>18</sup>In this version of the paper, banks portfolios are treated as static. In this setting a fully hedged factor is one for which a static hedge ensures the bank has no exposure. For example, if banks delta-hedge a factor to which they have a nonlinear exposure, they are non-linearly exposed to the factor. Assuming the non-linearity is captured by the exposure measures, using the methods in this paper the factor could in theory be identified as important for systemic risk purposes, and would not be considered unhedged.

<sup>19</sup>The assumption that all banks hedge some factors  $F_B$  is made to illustrate ideas regarding what factors SIR can detect. The more general point is banks have more exposure to some factors than to others. The factors to which banks as a group have more exposures are more important for systemic risk. The method in this paper differentiates among the factors it detects based on their importance for systemic risk. Stress tests are constructed based on those factors. I thank without implicating Rama Cont for encouraging me to clarify this aspect of the paper.

$\epsilon_i$ . Mathematically, this implies  $V_i(X)$ , the value of bank  $i$ 's portfolio as a function of  $X$ , reduces to a function of  $F_A$  and  $\epsilon_i$ ,

$$V_i(X) = V_i(F_A, \epsilon_i), \quad (2)$$

which stacked across banks has form

$$V(X) = V(F_A, \epsilon). \quad (3)$$

Equations (3) and (1) are useful for thinking about the types of risks that systemic stress-tests may be designed to control. One form of risk is that banks exposures to common unhedged factors as captured by  $F_A$  in equation (3) cause them to become financially impaired at the same time. A second type of systemic risk is that banks hedging strategies may fail for some reason such as a counterparty default in which the party that is providing a hedge against factor such as  $F_B$  cannot provide it when required to do so.<sup>20</sup> This paper is primarily focused on systemic risk due to common factor exposures.<sup>21</sup> Systemic impairment is written as  $SysImpair(V(X))$  because it depends on  $V$ , the vector of the values of banks net worth. Systemic impairment will also be written as  $SysImpair(F_A, \epsilon)$  to emphasize its dependence on common factors and residual risks.

The relevant factors for creating stress tests are  $F_A$ . The challenge is how to “identify” those factors, where formally identification of the factors means the identification of the space spanned by the factors.<sup>22</sup>

To identify the factors, I assume an approximation of the mapping between the variables  $X$  and the value of each banks portfolio  $V_i(\cdot)$  is known, or knowable to regulatory authorities:

**Assumption 2** *Regulators have approximations of  $V_i(X)$  that are sufficient to identify the factors.*

The assumption that regulators have approximations of  $V_i(X)$  is increasingly realistic. For example, in the case of stress-tests for market risk, the Federal Reserve collects risk sensitivities for approximately 30,000  $X$  variables, where each sensitivity measures how the value of the portfolio changes for small to medium-size changes in individual  $X$  variables. Similarly, for positions in the banking book, the Federal Reserve receives detailed information on banks loan portfolios, including for example information on each wholesale C&I loan that has value of at least 1 million dollars.

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<sup>20</sup>This second type of systemic risk is in the process of being addressed through policy reforms that control counterparty credit risk. These reforms include the migration of bilateral derivatives positions to CCPs and higher margin requirements on bilateral derivatives trades.

<sup>21</sup>Although not the paper’s main focus, as discussed further below the methodology in the paper can be used to identify hedged exposures for which further systemic risk analysis may be warranted.

<sup>22</sup>If  $\psi$  is  $K_A \times K_A$  and has full rank, then  $\psi F_A$  and  $F_A$  contain same statistical information about  $V(\cdot)$ . Therefore, the factors  $F_A$  can only be identified up to a rotation matrix  $\psi$ .

This banking book information is used to analyze how movements in economic variables are likely to affect the value of the loans. If the value of the approximations depends on the factors  $F_A$ , then under additional regularity conditions discussed below, it is also likely that the factors will be identifiable, as discussed further below.

The steps used to identify the factors are the following:

1. Draw a realization of the vector  $X$  from its distribution.
2. Compute  $V(X)$ .
3. Compute  $SysImpair[V(X)]$
4. Repeat steps 1-3  $Ndraws$  times.<sup>23</sup>
5. Use the simulated values of  $SysImpair(V(X))$  and  $X$  in Sliced Inverse Regression (SIR) factor analysis to identify the space spanned by the factors  $F_A$ .

Intuition for why this approach can identify the factors  $F_A$  comes from the steps. In step 1,  $X$  depends on  $F_A$ ,  $F_B$ , and  $U$ . In step 2, because  $F_B$  is hedged,  $V(X)$  only depends on  $F_A$  and  $\epsilon$ . Therefore, in step 3, systemic impairment is only a function of  $F_A$  and  $\epsilon$ :  $SysImpair[V(X)] = SysImpair(F_A, \epsilon)$ . In step 5, sliced inverse regression projects the simulated values of the  $X$  variables onto the simulated values of  $SysImpair(F_A, \epsilon)$ . If the  $X$  variables are independent of banks remaining idiosyncratic risk  $\epsilon$ , then the projected values of the  $X$  variables,  $E[X|SysImpair(F_A, \epsilon)]$  will only be functions of  $F_A$ . Under certain regularity conditions described below, it will then be possible to use the projections to identify factors that lie within a subspace of the space spanned by  $F_A$ ; under some conditions the identified factors will span the same space as  $F_A$ . Moreover, the identified factors will turn out to be principal components that are ranked by their ability to explain systemic impairment. Because the relationship between the principal component factors and the  $X$  variables can be estimated, changes in the factors can be used to find the size and direction of movements in the  $X$  variables that are most likely to contribute to the risk of systemic impairment.

Steps 1-4 provide an ideal setting to apply SIR in step 5. For step 5, the following assumptions are made to identify the space spanned by factors:

**Assumption 3** *There are  $K_A$  factors  $F_A$  that affect systemic impairment. Each of the  $K_A$  factors  $F_{A,k}$  is expressible as a linear combination of the  $X$  variables.*

$$F_{A,k} = X\beta_k, \quad k = 1, \dots, K, \quad (4)$$

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<sup>23</sup>The draws of  $X$  should be made from the conditional distribution of  $X$  viewed as appropriate for the stress-test exercise. For example, if a goal is to capture stress due to high energy prices, then draws of  $X$  should come from their distribution conditional on high energy prices. The draws of  $X$  should be i.i.d.

where each of the  $\beta_k$  vectors is  $N \times 1$ .

**Assumption 4** *The  $X$  variables are distributed independently of the vector of banks residual risks  $\epsilon$ .*

**Assumption 5** *For every  $N \times 1$  vector  $b$ , there exist constants  $c_k(b)$ ,  $k = 0, \dots, K_A$  such that*

$$E(Xb|X\beta_1, \dots, X\beta_K) = c_0(b) + \sum_{k=1}^K c_k(b)X\beta_k \quad (5)$$

Assumption 3 is equivalent to assuming that the information in the factors that generate systemic risk are expressible as  $K_A$  linear combinations of the  $X$  variables that affect the banks.<sup>24</sup> It therefore follows that systemic impairment has the functional form

$$SysImpair(F_A, \epsilon) = SysImpair(X\beta_1, X\beta_2, \dots, X\beta_{K_A}, \epsilon). \quad (6)$$

Assumption 4 has the implication that idiosyncratic risk at the portfolio level ( $\epsilon_i$ ) for each bank  $i$  cannot be used to forecast the variables  $X$ . This assumption cannot literally be true because  $\epsilon$  for each bank depends on the idiosyncratic risk of the variables  $X$ , but the assumption holds approximately since the forecasting power of the residuals approaches zero in diversified portfolios.<sup>25</sup> Put differently, assumption 4 should be interpreted as an assumption that the large banks to which stress-testing is applied hold diversified portfolios.

Assumption 5 states that the expected value of linear combinations of the  $X$  variables given the systemic risk factors is a linear combination of the systemic risk factors. This assumption will be satisfied if the  $X$  variables are elliptically distributed. As discussed in Li (1991), the methodology for uncovering the factors also works well even if this assumption holds approximately.

In step 5, the systemic risk factors are identified using the Sliced Inverse Regression (SIR) method of Li (1991) as refined using the Correlation Pursuit (COP) methodology of Zhong et al (2012). The main intuition for how SIR identifies the space spanned by the systemic risk factors will be presented in this subsection. Further information on SIR and COP is presented in the appendix.

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<sup>24</sup>This is similar in spirit to factor-mimicking portfolios that are often used in empirical asset pricing studies.

<sup>25</sup>For example, suppose  $X = f + U$ , where  $U$  is i.i.d., and bank  $i$ 's portfolio has exposure of  $1/N$  to each of the  $X$  variables. Then the portfolio's return,  $R_i$  is given by  $R_i = f + (1/N) \sum U_i$ , where the term in parenthesis is  $\epsilon_i$ . The covariance between any element of  $X$  such as  $X_j$  and  $\epsilon_i$  is  $(1/N)\sigma^2(U)$ , which vanishes with  $N$ , showing that the residual return of the portfolio has very little power to forecast  $X_j$ . By contrast the covariance between  $X_j$  and the systematic component of the portfolio returns is  $\sigma^2(f)$ , which does not vanish with  $N$ . This shows the portfolio's return has power to forecast the elements of  $X$  because of the portfolio's exposure to the factor risk; the idiosyncratic part of the portfolio's return, by contrast, has essentially no forecasting power.

SIR relies on inverse regression in which each of the simulated  $X$  variables is nonparametrically regressed on the simulated measure of systemic impairment  $SysImpair(F_A, \epsilon)$ , to compute the fitted value  $E[X|SysImpair(F_A, \epsilon)]$ . By assumption 4, the fitted value does not depend on the banks portfolios' idiosyncratic risk  $\epsilon$ , it only depends on  $F_A$ . To recover the space spanned by the factors  $F_A$ , *SIR* performs a principal components analysis based on the fitted values.

To economize on notation below,  $SysImpair(F_A, \epsilon)$  will be denoted  $Y(X\beta_1, \dots, X\beta_{K_A}, \epsilon)$ , or simply as  $Y$ . The factors  $F_A$  will be used interchangeably with  $X\beta_1, \dots, X\beta_K$ .  $\Sigma_{XX}$  denotes the variance covariance matrix of  $X$  and  $\Sigma_{E(X|Y)}$  denote the variance covariance matrix of the fitted values:

$$\begin{aligned}\Sigma_{XX} &= E\{[X - E(X)]' \times [X - E(X)]\} \\ \Sigma_{E(X|Y)} &= E\{[E(X|Y) - E(X)]' \times [E(X|Y) - E(X)]\}\end{aligned}$$

Sliced Inverse Regression identifies a subspace of the space spanned by the factors as the vectors  $Xb_k$  where the  $b_k$  vectors are solutions to the problem:

$$\text{Max}_{b_k} b_k' \Sigma_{E(X|Y)} b_k \tag{7}$$

subject to the constraint

$$b_k' \Sigma_{XX} b_k = 1,$$

and subject to the condition that the  $b_k$  vectors are orthogonal  $b_k' b_j = 0$  for  $j \neq k$ .

When SIR is used to estimate the  $b_k$  coefficients, it does so using sample estimates of  $\Sigma_{XX}$  and  $\Sigma_{E(X|Y)}$ . The analysis in this section illustrates the information that *SIR* recovers about the factors when  $\Sigma_{XX}$  and  $\Sigma_{E(X|Y)}$  are known. Distribution theory for the  $b_k$  coefficients is contained in Li(1991), Chen and Li(1998), and Zhong et al (2012).

The first order condition for choosing  $b_k$  is:

$$\Sigma_{E(X|Y)} b_k = \lambda_k \Sigma_{XX} b_k,$$

where  $\lambda_k$  is the Lagrange multiplier on the constraint. Rearrangement shows  $b_k$  and  $\lambda_k$  are eigenvectors and eigenvalues of  $\Sigma_{XX}^{-1} \Sigma_{E(X|Y)}$ :

$$\Sigma_{XX}^{-1} \Sigma_{E(X|Y)} b_k = \lambda_k b_k, \tag{8}$$

and that the  $Xb_k$  are therefore principal components constructed from  $\Sigma_{XX}^{-1}\Sigma_{E(X|Y)}$ . Because the  $b_k$  coefficients are eigenvectors, they are orthogonal, and thus the orthogonality condition does not constrain them. Following Zhong et al (2012), each  $b_k$  vector is referred to as a principal direction. The number of principal directions is the number of statistically significant eigenvalues from equation (8).

The principal components are not the systemic factors, but they lie within a subspace of the space spanned by the factors. When the number of principal directions is equal to the number of factors, then the principal components and the factors  $F_A$  span the same space. An advantage of focusing on the principal directions for modeling systemic impairment is that the eigenvalues measure the principle components based on their ability to statistically explain systemic impairment, with the larger eigenvalues corresponding to more explanatory power.<sup>26</sup>

The proposition and corollary that follow show that the principal components are spanned by the factors, and when both have the same dimension they span the same space.

To illustrate that the  $Xb_k$  vectors that are identified in the maximization problem (7) are spanned by the factors, note that any  $Xb_k$  can be decomposed into its projection on the factors ( $= \sum_{k=1}^K c_k X\beta_k$ ) and into a component  $Xb^\perp$  that is orthogonal to the factors. Because the projection component is spanned by the factors, it suffices to show that  $b$  vectors that solve equation 7 cannot contain an orthogonal component  $b^\perp$ . The theorem and proof of this result is based on Li (1991).

**Proposition 1** *For the  $b_k$  coefficients that satisfy equation 8, each principal component  $Xb_k$  is spanned by the factors  $X\beta_k$ ,  $k = 1, \dots, K_A$ .*

**Proof:** See the appendix.  $\square$ .

The main step in the proof shows that  $\text{Var}[E(Xb^\perp|Y)] = 0$  ( $= b^{\perp'}\Sigma_{E(X|Y)}b^\perp = 0$ ), or equivalently, that  $E(Xb^\perp|Y)$  is a constant that does not vary with  $Y$ .<sup>27</sup> To see that it is a constant, note that the information contained in  $Y$  is  $\epsilon$  and the factors  $X\beta_k$ ,  $k = 1, \dots, K_A$ . By assumption 4, the  $\epsilon$  coefficients have no power for forecasting  $E(Xb^\perp)$ . By assumption 5,  $E(Xb^\perp|X\beta_k, k = 1, \dots, K_A)$  is linear in the  $X\beta_k$ , but also by definition  $Xb^\perp$  is uncorrelated with each of the  $X\beta_k$ . It follows that  $E(Xb^\perp)$  does not change with the  $X\beta_k$ , and therefore that  $E(Xb^\perp|Y)$  does not vary with  $Y$ . This means any  $Xb_k$  that solves equation (8) is spanned by the factors.

**Corollary 1** *If the number of principal directions is equal to the number of factors, then the factors and the principal components span the same space.*

<sup>26</sup>See appendix B for details.

<sup>27</sup>If  $E(Xb^\perp)$  is a constant that does not vary with  $Y$ , then it follows that  $\Sigma_{E(X|Y)}b^\perp = 0$ . The proof then show that any  $b_k$  that solves equation (8) must have as its  $b^\perp$  component  $b^\perp = 0$ .

**Proof:** Let  $B$  denote the matrix  $(b_1, b_2, \dots, b_{K_A})$  and  $\beta$  denote the matrix  $\beta_1, \dots, \beta_{K_A}$ . Since  $b$  and  $\beta$  are non-singular and  $\beta$  spans the elements of  $B$ ,  $B = \beta\Pi$  for some non-singular  $\Pi$ . Therefore  $\beta = B\Pi^{-1}$ , and therefore  $\beta$  is also spanned by  $B$  and both span the same space.  $\square$ .

The corollary shows that *SIR* will identify the space spanned by the factors provided that the rank of  $\Sigma_{XX}^{-1}\Sigma_{E(X|Y)}$  has the same rank as the number of factors. Although *SIR* can be used to identify large parts of the factor space that are important to systemic risk, it is important to emphasize it can fail to identify factors in cases when  $E(X|Y)$  does not change with  $Y$ , even though a factor affects  $Y$ . For example, if the factor is just  $X_1$  and  $Y$  is a symmetric function of  $X_1$  such as  $Y = bX_1^2$ , and  $X_1$  is standard normal, then  $E(X_1|Y) = 0$  and therefore *SIR* could not detect  $X_1$  as a factor in this simple example.

When using *SIR*, identification of the space spanned by the factors relies on estimates of the matrix  $\Sigma_{XX}^{-1}\Sigma_{E(X|Y)}$ . When  $X$  is high dimensional and the time series on  $X$  is short, then estimates of  $\Sigma_{XX}$  and its inverse are likely to be inaccurate. As noted in Zhong et al (2012), this problem, if not addressed, will reduce the accuracy of *SIR* when the number of potential  $X$  variables is large. The Correlation Pursuit (*COP*) methodology of Zhong et al (2012) is designed to address this difficulty. The underlying assumption in Zhong et al is that a relatively sparse subset of the  $X$  variables, denoted  $x$ , contains the essential information on the factors. Conditional on  $x$  the information on  $Y$  contained in the other  $X$  variables is assumed to be redundant.<sup>28</sup> The *COP* methodology chooses the relevant variables  $x$  based on their ability to create principal components that explain  $Y$ . For this paper, the elements of  $x$  are chosen based on their ability to explain systemic impairment. *COP* selects the relevant  $x$  variables by starting with a candidate set of active variables  $x_0 \in X$ ; it then scrolls through the remaining variables in  $X$  and performs variable addition and deletion steps that add (delete) variables to (from) the active set if they statistically improve (don't improve) explanatory power for  $Y$ . Zhong et al show that under a set of regularity conditions as the size of the sample of  $X$  and  $Y$  variables approaches infinity, *COP* consistently chooses the set of  $x$  variables that are relevant for determining the factors that explain  $Y$ . Although Zhong et al provide asymptotic theory for choosing  $x$  consistently, they emphasize that the asymptotic theory treats  $K_A$  as known when it is not, and also the asymptotic theory for adding and deleting variables makes strong assumptions that can only hold as the sample size approaches infinity. Therefore, for finite samples they recommend choosing  $K_A$  based on the BIC criterion; and they choose the critical values for determining whether to add or delete variables using cross-validation. Further details on how to implement *SIR* and *COP* are provided in the appendix.

In summary, this subsection has illustrated an approach using *SIR* and *COP* to identify the relevant factors that explain systemic impairment given how banks hedge and it has presented an approach for identifying the variables  $x$  that explain these factors. The next section provides details on how to use the identified variables and factors to create stress tests for systemic risk.

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<sup>28</sup>This is a reasonable assumption in many circumstances. For example, much of the information in the term structure of interest rates can be summarized by the information in a small number of yields.

### 3.2 Choosing Stress Scenarios for Systemic Risk

This section uses the principal components extracted in the last subsection to create stress scenarios and resulting capital injections to ensure the financial system is resilient against systemic risk with high probability. To implement the methodology, three elements are required. First, a measure of systemic impairment is needed to serve as the  $Y$  variable in the last section, as well as to measure systemic risk. Second, a method is needed to define stress scenarios in terms of movements in the systemic risk factors. Third, a method is needed to choose stress scenarios such that if banks are well capitalized against the scenarios considered, then they will be well capitalized with high probability against systemic impairment; i.e. systemic risk will be low. Below, each of these elements is provided in turn.

#### Measuring systemic impairment

Recall that systemic impairment is the event that too many banks become financially distressed together, and hence cannot provide needed financial intermediation services to the real sector. Financial distress is measured based on banks equity capital (net worth) relative to its risk. For example, if the stress horizon is one year, and at the end of that year a banks capital is low, while the volatility of capital is high, then the bank is likely to become insolvent shortly thereafter. It will therefore not be able to raise funds to intermediate loans, and hence its financial distress will be high.

The analysis on systemic impairment measurement is based on Pritsker (2013). Without loss of generality, banks are stress-tested at date 0, and the stress-test horizon is normalized to be one period. There are  $J$  financial intermediaries  $j = 1, \dots, J$ . At date 0, each financial intermediary has equity  $E_j$ , and liabilities  $L_j$  that finance assets  $A_j (= E_j + L_j$  by the balance sheet identity). Bank  $j$ 's asset portfolio has return  $R_j$  between date 0 and date 1. Additionally, the gross return earned by its liability holders is  $\bar{R}_{l,j}$ . Thus, bank  $j$ 's capital ratio at date 0 is  $C_j(0) = \frac{E_j(0)}{A_j(0)}$ , and its capital ratio at date 1 is  $C_j(1) = \text{Max}[\frac{E_j(1)}{A_j(1)}, 0]$ . The capital ratio for bank  $j$  at date 1 can be written as a function of its initial capital ratio and the return on its assets:

$$\begin{aligned}
 C_j(1) &= \text{Max}\left(\frac{E_j(1)}{A_j(1)}, 0\right) \\
 &= \text{Max}\left(\frac{A_j(1) - L_j(1)}{A_j(1)}, 0\right) \\
 &= \text{Max}\left(1 - \frac{L_j(0)\bar{R}_{l,j}}{A_j(0)R_j}, 0\right) \\
 &= \text{Max}\left(1 - \frac{(1 - C_j(0))\bar{R}_{l,j}}{R_j}, 0\right)
 \end{aligned}$$



As a result of the stress-test conducted at date 0, banks may be required to inject more equity into the bank. I assume this additional equity is invested at the risk free rate, and earns a gross return of  $R_f$ . If the equity injected at date 0 is equal to a fraction  $CI_j$  of initial assets, then assets at date 1 become  $A_j(0)R_j + A_j(0)CI_jR_f$ . Making this substitution, bank  $j$ 's capital ratio at date 1 is given by

$$C_j(1) = \text{Max} \left( 1 - \frac{[1 - C_j(0)]\bar{R}_{l,j}}{R_j + CI_jR_f}, 0 \right) \quad (9)$$

The volatility of bank  $j$ 's date 1 capital ratio as of date 1 is denoted  $\sigma(C_j(1))$ . Bank  $j$ 's financial distress at date 1 is modeled as a decreasing function of its capital ratio normalized by its volatility:  $D_j(\frac{C_j(1)}{\sigma(C_j(1))})$ . This ratio is inversely related to default likelihood after period 1, and therefore distress goes down as the ratio goes up. For convenience distress is parameterized to lie between zero and one ( $D_j(\cdot) \in [0, 1]$ ).

To model systemic impairment, I make the following assumptions:

**Assumption 6** 1. Each banks maximal financial intermediation capacity is proportional to its assets:  $FICapacity(j) = \gamma A_j$  with a constant of proportionality  $\gamma$  that is the same for all banks.

2. The fraction of a bank's maximal intermediation capacity that is lost in a scenario is proportional to its distress in that scenario:

$$\text{Loss of } j\text{'s capacity} = D_j \left( \frac{C_j(1)}{\sigma(C_j(1))} \right) \gamma A_j.$$

3. Systemic impairment occurs when the fraction of the economy's maximal intermediation capacity that is lost exceeds a threshold  $\zeta$ .

These assumptions capture the ideas the larger banks, measured by the size of their balance sheets, have more intermediation capacity, and that therefore more intermediation capacity is lost when larger banks are more financially distressed. It is assumed that when a little intermediation capacity is lost, other banks can step in and fill the capacity that is lost. But, when too much maximal capacity is lost, it becomes too large for others to fill in, resulting in systemic impairment.

Under assumption 6, the fraction of maximal intermediation capacity that is lost given a realization of banks return vector  $R_1, \dots, R_J$ , and given the Capital Injections received by banks, is denoted System Assets in Distress, abbreviated SAD:

$$SAD(R_1, \dots, R_J, CI_1, \dots, CI_J) = \frac{\sum_{j=1}^J D_j(\frac{C_j(1)}{\sigma(C_j(1))}) \gamma A_j}{\sum_{j=1}^J \gamma A_j} = \sum_{j=1}^J w_j D_j(\cdot) \quad (10)$$

Note, that the arguments of the capital ratios that are made explicit on the left hand side of the expression for SAD are for simplicity suppressed on the right hand side, and will be suppressed whenever it is convenient to do so.

The constant of proportionality  $\gamma$  drops out of the expression for  $SAD$ . As a result, it reduces to a weighed average of each banks distress function where each banks weight is its assets as a fraction of all banks assets.

Systemic risk is a function of the distribution function of systemic impairment. For simplicity, in this paper systemic risk, denoted  $\psi(0, T)$  is defined as the probability that systemic impairment occurs at the end of the time horizon  $T$  of the stress-test:

$$\psi = \text{Prob}[SAD(T) \geq \zeta].$$

## Creating Stress Scenarios

A stress scenario specifies values for all of the  $X$  variables. As noted in the introduction, choosing an appropriate scenario to achieve regulatory objectives is difficult when  $X$  is high dimensional. To reduce dimensionality, this paper defines stress-scenarios in terms of the identified factors, and then sets the  $X$  variables in the scenario to their expected values given the factors.

**Definition 1** *A Systemically Chosen Stress Scenario is a specification of realizations for the systemic risk factors  $F_A$  and a specification for the expected realizations of the other relevant  $X$  variables for determining the value of banks conditional on  $F_A$ .*

To compute the expected value of the  $X$  variables, consistent with assumption 5, the  $X$  variables are modeled as a linear function of the factors, and take a form that can be estimated by OLS through regressing the simulated values of the  $X$  variables on the simulated value of the factors<sup>29,30</sup>  
:

$$X_i = \alpha_i + F_A \theta_i + \epsilon_i \tag{11}$$

where  $\alpha_i$  is the  $N \times 1$  regression intercept,  $\theta_i$  is  $K_A \times 1$  vector of regression coefficients and  $\epsilon_i$  is an  $N \times 1$  residual.

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<sup>29</sup> Assumption 5 implies  $E(X|F_A)$  is a linear function of  $F_A$ . This is consistent with the OLS regression specification in equation 11.

<sup>30</sup> Recall the identified risk factors are  $F_{A,k} = xb_k, k = 1, \dots, K_A$ . Because the  $X$  variables and the  $x$  variables are simulated and the  $b_k$ 's are estimated, the simulated risk factors are "observable" in the simulation, as are the  $X$  variables. This makes it possible to estimate the relationship between the  $X$  variables and the risk factors.

Using the definition and equation (11), if the chosen factor realizations are  $\tilde{F}_A$ , then the stress-scenario is given by:

$$X_i = \alpha_i + \tilde{F}_A \theta_i, \quad X_i \in \mathbf{X} \quad (12)$$

To assure that banks have enough capital on the basis of a stress test, it is also necessary to know how each banks financial distress is related to the factors. To model this, for simplicity I assume banks liabilities are unaffected by stress, that each bank's asset returns are linearly related to the  $X$  variables with  $\beta_{i,j}$  representing the sensitivity of  $R_j$  to  $X_i$ , and that the realizations of the  $X$  variables fully explain banks returns.<sup>31</sup> With this formulation, for each bank  $j$ ,  $R_j$  can be expressed as a linear function of the systemic factors, and a residual term that will be correlated across banks because many banks are exposed to common  $\mathbf{X}$  variables:

$$\begin{aligned} R_j &= \alpha_{0,j} + \sum_i \beta_{i,j} X_i \\ &= \alpha_{0,j} + \sum_i \beta_{i,j} (\alpha_i + F_A \theta_i + \epsilon_i) \\ &= \alpha_j + F_A \theta_j + \epsilon_j, \end{aligned} \quad (13)$$

where the  $\beta_{i,j}$  coefficients are scalars and

$$\begin{aligned} \alpha_j &= \alpha_{0,j} + \sum_i \beta_{i,j} \alpha_i \\ \theta_j &= \sum_i \beta_{i,j} \theta_i \\ \epsilon_j &= \sum_i \beta_{i,j} \epsilon_i \end{aligned}$$

## Formulating Stress Scenarios based on a Systemic Risk Objective

The main result in the paper is if  $SAD$  is approximated by  $ASAD$ , a variant of  $SAD$  that linearizes the relationship between  $SAD$  and  $R_j + CI_j R_f$ , then if regulators objective function is defined in terms of  $ASAD$ , then there is a stress-scenario and resulting capital injections that assures systemic risk is low. This is formally stated in the following proposition:

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<sup>31</sup>These assumptions can be relaxed to allow the liabilities to be affected by  $X$ , to allow the  $X$  variables to non-linearly affect banks asset returns and liabilities, and to allow the value of assets and liabilities to fluctuate for reasons other than the  $X$  variables. However, relaxing these assumptions significantly complicates the modeling.

**Proposition 2** *If SAD is linearly approximated by ASAD<sup>32</sup>:*

$$ASAD = C_0 + \sum_j D_{j,1}(R_j + CI_j R_f), \quad (14)$$

and the return on each bank  $j$ 's portfolio,  $R_j$ , satisfies equation (13), and if regulators systemic risk objective is to ensure that

$$Prob(ASAD \geq \xi) \leq \psi,$$

then there is systemic risk factor shock  $F_A^*$  such that when the stress scenario is  $X_i = \alpha_i + F_A^* \theta_i$  for all  $X_i$ , and banks inject capital equal to the present value of their losses in the stress scenario, then after the capital is injected,  $Prob(ASAD \geq \xi) \leq \psi$ .

**Proof:** See the appendix.

To provide intuition for the proposition note that equation (14) for  $ASAD$  and equation (13) together imply that  $ASAD$  has a linear stochastic component that depends on the factors  $F_A$  and non-factor risks, and a linear component in terms of the capital injected by banks:

$$ASAD = C_0 + \sum_j D_{j,1}(\alpha_j + F_A \theta_j + \epsilon_j + CI_j R_f) \quad (15)$$

$$= C_0 + \alpha + F_A \theta + \epsilon + CIE, \quad (16)$$

where  $D_{j,1}$  is negative since more capital reduces banks financial distress,  $\alpha = \sum_j D_{j,1} \alpha_j$ , and other terms in the final equation defined similarly.

The expression shows the magnitude of  $ASAD$  can be controlled by capital injections, summarized by the  $CIE$  term. The first part of the proof finds the least negative value of  $CIE$ , denoted  $CIE^*$ , that just satisfies regulators objective for systemic risk.<sup>33</sup>

$$CIE^* = CIE : Prob(C_0 + \alpha + F_A \theta + \epsilon + CIE \geq \xi) = \psi.$$

The second part of the proof finds values of the systemic factor  $F_A^*$  that satisfy the condition that if the stress-scenario is

$$X_i = \alpha_i + F_A^* \theta_i$$

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<sup>32</sup>Some arguments of the linearization are suppressed for simplicity, including  $(C_j(1))$  which is assumed fixed. The parameters of the linear approximation to  $SAD$ , the scalar  $C_0$  and the scalars  $D_{j,1}$  ( $j = 1, \dots, J$ ), can be derived from a first-order Taylor series for  $SAD$ , or they can be estimated by linearly regressing simulated values for  $SAD$  on simulated values of  $(R_j + CI_j R_f)$ .

<sup>33</sup>Solving for  $CIE^*$  requires knowledge of the CDF of  $F_A \theta + \epsilon$ , denoted  $H(\cdot)$ . Finding this CDF is relatively straightforward because although  $H(\cdot)$  is not known, it is relatively easily estimated since  $F_A \theta + \epsilon$  is a single random variable, simulated values of  $F_A$  are available from  $COP/SIR$ , and simulated values of  $\epsilon$  can be constructed using estimated versions of equations (11), and (13), with the  $D_j$  coefficients from equation (14).

for all  $i$ , then if banks inject enough equity capital to cover their losses (measured from their net returns), then the resulting capital injections ensure  $CIE = CIE^*$ , thus achieving the systemic risk objective.<sup>34</sup> The condition for  $F_A^*$  is the equation

$$F_A^* \theta = -CIE^* - \alpha + \sum_j D_j. \quad (17)$$

When there is only one systemic risk factor, then  $\theta$  is a nonzero scalar, and this equation has one solution for  $F_A^*$ . When there is more than one factor, then  $\theta$  is a vector, and there are multiple solutions for  $F_A^*$ , which means there is room to choose  $F_A^*$  to satisfy equation (17), while also satisfying other side criteria. Two criteria are considered here, maximum likelihood and minimum cost.

The maximum likelihood criterion chooses the value of  $F_A^*$  to satisfy equation (17) and have maximal likelihood. To solve for the maximum likelihood  $F_A^*$ , note that  $F_A$  has mean 0 since each element of  $X$  is normalized to have mean 0, and  $F_A$  has variance  $I$  since  $F_A$  is a matrix of principal component factors. Under the auxiliary assumption that  $F_A$  is also multivariate Gaussian, then it is straightforward to show that the maximum likelihood value for  $F_A^*$ , denoted  $F_A^*(Maxlik)$ , is:

$$F_A^*(Maxlik) = \frac{(-CIE^* - \alpha + \sum_j D_j) \theta'}{\theta' \theta}. \quad (18)$$

An alternative criterion for choosing  $F_A^*$  is to find the stress scenario that minimizes banks capital costs while at the same time satisfying the constraint in equation (17).  $F_A^*$  is chosen inefficiently if the ensuing capital requirements from the stress scenario inject large amounts of capital into banks for which the marginal systemic risk benefits are small, while injecting too little capital in banks with high marginal benefits. The following problem for choosing  $F_A$  minimizes the costs of injecting capital while choosing a stress scenario that satisfies the constraint in (17), where  $\lambda$  is the marginal cost of injecting equity capital.

$$F_A^*(ef fic) = Argmin_{F_A} - \lambda \sum_j A_j \text{Min}(-1 + F_A' \theta_j + \alpha_j, 0) \quad (19)$$

such that

$$\sum_j D_{j,1} \text{Min}(-1 + F_A' \theta_j + \alpha_j, 0) = -CIE^*$$

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<sup>34</sup>Banks capital injections equal the present value of their losses in the stress-scenario discounted at the risk-free rate. Since the capital is assumed to be invested risk-free, it produces enough capital to cover losses at date 1 in the stress scenario.

Note, in this optimization problem the “Min” operator on the right hand side of equation (19) rules out negative capital injections. These are ruled out because if they are allowed, then the optimization problem becomes a linear objective function with linear constraints, and thus would not have a bounded solution.

To the best of my knowledge, this is one of very few papers (the only ?) that have derived stress-tests with the explicit goal that the resulting capital injections satisfy an explicit systemic risk objective, and that moreover are designed to guarantee that if the banking system is well capitalized against the scenario, it is well capitalized against other plausible scenarios with a high likelihood that is chosen by the regulator.

Several qualifiers regarding proposition 2 are important. First, because *ASAD* approximates systemic impairment, satisfying regulators systemic risk objectives based on *ASAD* will not necessarily satisfy the objectives based on *SAD*. This suggests using the *ASAD* approximation with side conditions (equations (18) and (19) ) to find the direction in which to move the systemic risk factor vector. Then, the amount by which the factors are moved in the appropriate direction is chosen until the resulting stress scenario achieves a systemic risk objective based on *SAD*.<sup>35</sup> As shown in the next section, this often works well in practice.

Second, there are circumstances when the capital injections required to achieve a systemic risk objective are unattainable on the basis of a single stress scenario. As a simple illustration of such circumstances, if banks only asset holdings are positions in a stock index, then the index is the factor  $F_A$ . If half the banks have identical long positions in the factor, while the other half have identical short positions, then absent capital injections, too much systemic impairment can occur if the factor takes very high values or very low values. This example is an extreme failure in which *ASAD* is approximated as a linear function of the factor but *SAD* is strongly non-linear. The consequence of the failure is that a single stress scenario in which the index drops can ensure that half the banks are well capitalized against that scenario, but, it cannot ensure that the other half of the banks are also well enough capitalized against stocks rising. Hence, in the example a single stress scenario is not sufficient to attain the objective. In this example, the solution is two scenarios, one in which stocks fall significantly and one in which they rise significantly. An advantage of the SCSA approach advocated in this paper is that part of choosing the stress scenario entails checking whether the regulatory objective is attainable with a single scenario, or whether more scenarios are needed.

In addition to these qualifications, there are areas where there is scope to further extend the SCSA approach. These include:

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<sup>35</sup>For example, if there are two systemic risk factors  $F_A(1)$  and  $F_A(2)$ , the chosen direction might specify that  $F_A(1)$  should increase 1.5 times as quickly as  $F_A(2)$ . Given this direction, increases in  $F_A(1)$  and  $F_A(2)$  can be solved for such that the resulting capital injections in the stress scenario achieve the systemic risk objective for *SAD*.

- Modeling the effects of interbank credit exposures.<sup>36,37</sup>
- Incorporating the modeling of banks liabilities and income as part of the analysis.
- Using the methodology for other measures of impairment, such as Systemic Expected Shortfall.

The next section of the paper illustrates the use of the SCSA methodology when applied to interest-rate positions.

## 4 SCSA applied to rates positions

In this section I analyze how well SCSA performs in generating systemic risk stress scenarios and ensuring banks are well capitalized against systemic risk. To do so, I apply SCSA for 6 hypothetical banks that for simplicity are modeled as only having trading book asset portfolios.<sup>38</sup> The assets of each bank are fixed income portfolios that are represented as positions in the zero coupon bonds of 8 countries (Australia, Canada, Germany, Japan, Sweden, Switzerland, the United Kingdom, and the United States).<sup>39</sup> A total of 83 zero coupon bonds are used in the analysis. The yield changes or the returns on these bonds are the  $X$  variables that are used to identify the factors for SCSA.<sup>40</sup> Because these bonds are for different countries, positions in these bonds have both interest-rate and foreign exchange risk. For simplicity, I assume below that the FX risk has been fully hedged.

The performance of SCSA is studied for simulated values of banks portfolio holdings according to the following steps:

1. Randomly create a portfolio of assets and liabilities for each of the six banks.
2. Use COP to choose the variables  $x \in X$  that are most important for systemic risk. Use SIR to identify the factors that are most important for systemic risk.
3. Analyze whether the factors identified by SIR are related to systemic risk.
4. Create stress-scenarios based on movements in the factors.
5. Analyze whether banks losses in the stress-scenarios are related to systemic risk.

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<sup>36</sup>The importance of accounting for interbank credit may be diminishing if CCP clearing requirements cut such risks enough, but could become very important if CCPs get into trouble.

<sup>37</sup>For potential directions to incorporate them See Ota (2013),Pritsker (2014).

<sup>38</sup>Six U.S. banks are subject to a market-risk stress test for their trading book positions as part of CCAR stress testing.

<sup>39</sup>For the United States the yield tenors in years are 1,2,3,4,5,7,8,10,12,15,20,25,30. For the other countries the tenors in years are 0.25,.5,.75,1,2,3,5,7,10,15.

<sup>40</sup>The results from using yield changes and bond returns are similar.

6. Repeat the previous steps  $M$  times.

The above steps are repeated a total of 20 times. The first 10 times these steps are performed, the portfolio weights for each bank in 1 are randomly generated and distributed i.i.d., with a bias towards having long holdings of zero coupon bonds. The second 10 times they are also i.i.d, but are not biased to be long or short.

Given the portfolios that have been generated for all 6 banks, in step 2 as part of doing COP and SIR it is necessary to simulate the  $X$  variables, compute the returns on banks portfolios, and compute SAD. For simplicity, SAD is computed in the trading book based on returns over a 1-month horizon. The returns for the zero coupon bonds over this horizon are simulated by bootstrapping from the historical time series of monthly returns. Further details on how banks portfolio weights and returns on the  $X$  variables are simulated are contained in the appendix, as is information on the functional form of the SAD objective function.

The results for the 10 sets of random portfolios (labeled Simulation 1 - Simulation 10 in the figures) which are long zero coupon bonds on average, show SAD and correlation pursuit usually identify a single systemic risk factor and uncover a very strong relationship between SAD and this systemic risk factor. In the case of Simulation 1, a scatter plot of SAD (labeled DV01 SAD, or DV01 for short) versus the factor shows a strong monotone relationship between SAD and the factor (Figure 1). The Correlation Pursuit and Sliced Inverse Regression Statistical procedures identify the factors (the space spanned by the factors), but do not identify the precise relationship between SAD and those factors. Nonparametric regression of SAD on the factors is utilized to explore the relationship between the expected value of SAD and the factors, and also to give an eyeball view of how much SAD deviates from its expected value conditional on the factors (Figure 2).<sup>41</sup> The regressions show that the relationship between expected SAD and the factor (the red curve labeled kernel) is slightly nonlinear. The figure also shows the nonparametric fit of the factor to SAD is very good.

While the nonparametric regression curve is suggestive of how much information is likely to be statistically captured by the factor, it does not establish that if banks hold enough capital to be resilient against a stress-scenario constructed based on movements in the factor, then it will ensure SAD is low with high probability. A necessary condition for the stress-scenario based on the factor to succeed in achieving capital adequacy is that the losses from stress scenarios based on the factor are highly correlated with SAD. Whether this correlation is high depends on how the factor is mapped to returns and then to SAD as part of the stress-testing procedure. Those mapping procedures are not accounted for as part of the nonparametric regressions; if the mapping procedures are not appropriate, then a stress-test based on the factor may fail to capture SAD well even if the factor is strongly statistically related to SAD. The extent to which mapping errors are

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<sup>41</sup>The regressions use Nadaraya-Watson kernel regression with a Gaussian kernel. The kernel bandwidth  $h$  was chosen as  $h = \sigma(f)N^{-1/5}$ , where  $N$  is the number of time-series observations of  $f$ .  $h$  was not optimally chosen.



an issue can be examined by stressing the factor, applying the mapping used in the stress test, and then comparing the mapped values to the nonparametric estimates and to the scatter plots. The results in the cases considered shows that SAD that results from stress tests based on the factor (labeled S.T. and plotted in green) track the nonparametric estimates of SAD fairly well (Figure 3) when the portfolio exposures are on average long. The results using all 10 simulated sets of long-on-average portfolios are similar (Figure 4).

The results for the 10 sets of portfolios appear to be very encouraging for finding systemic-risk factors that can be successfully used as the basis for a stress test. However, a more reasonable interpretation is that the method should work quite well for the case considered because in the case considered interest-rate factors should have significant explanatory power for SAD. To see why, note that for the method used to construct banks portfolio weights the expected DVO1s at each tenor on each yield curve are positive and the same. If banks portfolios are not too far from what is expected, their expected portfolios are sensitive to parallel yield curve shifts, which is known to be an important factor in yield curve modeling. Put differently, it is encouraging that the method appears to find a factor when it should. However, a more stringent test of the usefulness of the approach would be if banks portfolio compositions were more heterogeneous. The analysis of this case has only been done so far for portfolios that are on average neither long nor short. Intuitively, a portfolio that has long and short positions will have exposure to changes in the slope of the yield curve as well as changes in the level, and if portfolios differ in their mix of exposures to the two types of factors, then a one-factor model will not fit as well when modeling systemic impairment. This intuition is borne out from the simulations for all neither long nor short portfolios (Figure 5). Although the fit is worse, it is still very good.

Given that the factors and stress-scenarios based on the factors both track SAD fairly well, the work that remains is solving for stress-scenarios that achieve the systemic risk objective. This is done in the next subsection.

#### 4.1 Solving for CIE and stress-scenarios

CIE and the stress-scenarios were investigated in two settings. The first used linear approximation. By construction, the stress-scenarios and capital injections solved for based on linear approximation satisfy the ASAD objective function exactly. It is appropriate to investigate whether they also satisfy the objective function of keeping SAD low with high probability. The answer in the cases investigated so far is that choosing stress scenarios using the expression for ASAD and equation (32) does not work well for SAD as parameterized because SAD is too nonlinear to be sufficiently well approximated along the length of its range. As an example for simulated portfolio 1, a QQ-plot of SAD versus ASAD shows that ASAD is well below SAD for high values of SAD (Figure 6). This implies that the capital injection required to reduce ASAD to the target level chosen by the

regulator can often be too small to reduce SAD by enough to achieve regulatory targets for SAD. In other words, the size of the stress scenario based on ASAD will tend to be too small, and therefore the size of the stress scenario needed to achieve the systemic risk objective needs to be solved for by other means. The approach I have pursued in the empirical analysis is just a slight modification of the linear approach with multiple factors. In particular, from equation (18),  $F^*$  is a scalar multiple of  $\theta$ , which is written below as:

$$F^*(MaxLik) = \kappa\Theta.$$

Let  $CI(\kappa)$  be banks required capital injections if the stress scenario is  $\kappa\Theta$ . Then I solve for  $\kappa$  such that if the required capital injections are  $CI(\kappa)$ , then the  $\text{Prob}(SAD(CI(\kappa)) \geq \zeta) \leq \psi$ . This step is not difficult since it simply involves simulating SAD with different levels of capital injections. For our preliminary analysis of 10 simulations of banks interest rate portfolio positions, this approach usually succeeded in generating a reasonable stress scenario and resulting capital injections that satisfied the systemic risk objective that  $P(SAD > .05) \leq .05$ . An example of the scenarios is provided in figure 7. In this case, only one systemic risk factor was identified, and hence changes in yields are proportional to that factor. In this case the movements in stress scenario can best be described as an upward shift in the yield curve that increases curvature.

Although in most of the portfolios considered a single reasonable stress scenario could ensure the banks were sufficiently capitalized to achieve the systemic risk objective, as shown in the next section, this is not always possible in general. Therefore, there are circumstances when more than one scenario will be required to achieve the systemic risk objective. A strength of the methodology above is it will fail to find a single satisfactory scenario when more than one is needed. In such circumstances, it becomes necessary to solve for the best set of multiple scenarios. This topic is beyond the scope of the present paper.

## 4.2 Supplemental Analysis

The analysis above illustrates the potential for using SCSA to choose scenarios. The analysis in this subsection investigates three supplementary topics:

1. Out of Sample-Fit of Correlation Pursuit and Sliced Inverse Regression.
2. Additional evidence for the importance of choosing scenarios using exposure data.
3. How symmetry biases factor identification in SIR, and how to fix the bias.

## Out of Sample Fit

A potential issue with Correlation Pursuit and SIR is that it may choose variables and factors that over-fit the in-sample data and consequently fit poorly out-of sample. The out of sample fit could be poor because of over-fitting or because the distribution of the variables being simulated is mis-specified. The latter problem is not due to SIR or COP. To abstract from the latter problem, I assume that the true distribution of interest-rates in 10 countries are known, and follows the Dynamic Term Structure Model in Wright (2011) as modified in the appendix D. Given this process, 5,000 two-year sample paths of returns were generated, and banks portfolios were constructed to be perfectly correlated with one of the first three principal components of global yield curve changes. Then, using these portfolio weights, COP/SIR was used to identify the factors in-sample. Then, using additional observations from the same DGP, COP/SIR was used to again identify the factors in the out of sample data. If the data is severely over-fit in either sample, then the extracted factors in the two subsamples will not be highly correlated, but in fact they are pretty highly correlated, providing preliminary evidence that over-fitting does not appear to be a severe problem (Figure 8).<sup>42</sup> for the portfolios considered.

## Additional Evidence for the Importance of choosing scenarios using exposure data.

This subsection further examines the importance of using exposure data to choose stress-scenarios. Recall that section 2 presented special cases in which stress-scenarios that are not chosen while accounting for positions can miss substantial risks in banks' portfolios. This section illustrates the same underlying idea by examining how the factors that SCSA identifies as important in stress-testing change as portfolio composition changes. To examine this question, the analysis contrasts 6 banks with identical portfolios that only invest in the bond-market with 6 banks that are identical and invest 50% of their assets in the bond market and the rest in the stock market. For all of the analysis in this part, the return series for banks portfolios are modeled via historical simulation, in which it is assumed that the distribution of returns in the future on bonds and stocks is the same as was experienced based on past history.

The effect on the extracted factors is measured by how the factor shocks affect yield curves and stock returns. In the case of interest rate portfolio, SIR only extracts a single interest-rate factor. One standard deviation shocks to this factor primarily change the curvature of the yield curve for most of the 8 countries analyzed in the historical simulation analysis, suggesting that this particular randomly chosen set of 6 bond portfolios (one for each firm) is mostly exposed to a curvature factor (Figure 9). For the portfolio that represents a 50-50 mixture of stocks and zero coupon bonds, two factors are extracted. The most important factor (factor 1) steepens most yield curves and depresses stock market returns (Figure 10). The second factor, which is much

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<sup>42</sup>The out-of-sample data used 200, 500, or 1,000 observations.

less important, appears to capture yield curve curvature. These results together illustrate, not surprisingly that the important scenarios to consider for systemic risk purposes depend on banks portfolio composition, and therefore it is important to account for this composition when deriving stress scenarios to achieve systemic risk objectives.

## Symmetry and SIR

As illustrated in section 3.1, if the  $Y$  variables used in *SIR* are a symmetric function of the risk-factors, then *SIR* and *COP* may fail to detect the factors. To investigate this issue when  $Y$  is the *SAD* function, I created examples in which there are 6 banks and the banks have portfolios whose stochastic components are long or short the same portfolio of zero coupon bonds. In this setting the factor is the return on the portfolio, and if it happens that 3 banks are long and 3 banks are short, then *SAD* is a symmetric function of the factors. This is a situation when *SIR*/*COP* should have a difficult time identifying the factors. Figure 11 illustrates the problem. The figure is analogous to Figure 6 except that there is occasional symmetry. As a result, the single factor is sometime poorly identified. For example, in the second case of simulated banks, there are 3 long and 3 short banks, which is near perfect symmetry. As a result, the factor does a poor job of tracking *SAD*, and the stress-scenario created based on the factor also tracks *SAD* very poorly. Because *SAD* is constructed as a sum of each bank's distress function, it is possible to identify the factor by pre-analyzing banks individual distress functions. For example, if scatter plots of banks simulated distress functions against each other, show that some banks are highly positively correlated while others appear to be highly negatively correlated the factors can be identified by performing *SAD* and *COP* on the banks whose distress functions are positively correlated. Then the factor can be used to examine how well it tracks *SAD* across all banks, as well as how stress-scenarios based on the factor would track true *SAD*. Figure 12 performs this analysis for the banks in simulation 2 from Figure 11. The figure shows that this approach has potential to overcome some of the problems that symmetry can pose when using *SIR* and *COP*.

## 5 Extensions to account for portfolio uncertainty and network uncertainty

The preceding sections of the paper analyze whether banks are likely to jointly experience financial distress based on information that is known to regulators. Based on this information the paper designs stress-tests to ensure banks are well well capitalized against changes in the value of assets that cause too many banks to experience joint financial distress in both the scenarios considered as well as other scenarios.

However, the above analysis does not cover all sources of systemic risk. One source it does not cover is systemic risk from freezes in lending that are caused by uncertainty over exposures. For example, if the value of a particular asset class drops in value, and it is possible that some banks have large exposures to that asset class, but it is uncertain who is exposed, it could potentially lead to much higher spreads in lending markets, and a potential lending freeze. The lending freeze may itself cause systemic problems even though (due to the stress-testing advocated above or other reasons) banks are unlikely to experience joint distress directly from movements in the value of their assets. Similarly, if banks are potentially linked to other banks through interbank deposits or as derivatives counterparties, then if one bank experiences difficulties, banks that are perceived to be linked to it may be susceptible to funding difficulties.

The framework in this paper provides a way that stress testing can be used to partially address systemic risk stemming from uncertainty about positions or about interbank exposures: it can do so by requiring banks to hold more capital for the opacity of their positions to others. To give some feel for the main idea, recall that each bank  $j$ 's capital at date 1 (the stress horizon) is a function of the return on the bank's portfolio  $R_j$  between dates 0 and 1. This is a function of the movements in the variables  $X$ , and the banks direct risk exposures to  $X$ , which for simplicity are denoted  $\omega_{j,X}$ ; i.e.  $R_j = \omega_{j,X}X$ .

If banks risk exposures are not precisely known to the public, then when *SAD* is computed via simulation in section 3.2, then  $X$  and  $\omega_{j,X}$ ,  $j = 1, \dots, J$  should both be simulated from their priors based on public information, and in equation (13), returns should be computed accounting for variation in  $\omega_{j,X}$ . Using this treatment, banks simulated returns account for public uncertainty in their risk exposures and the consequent choice of the stress-scenario should automatically account for this uncertainty. I anticipate that this approach is likely to require banks that are more opaque to hold more capital—and that doing so will reduce the likelihood and severity of potential funding freezes due to uncertainty. Whether it actually can do so requires further investigation.

In addition to uncertainty about positions, the framework can also potentially be used to account for uncertainty about interbank exposures. To model this uncertainty, for now I focus on interbank deposits and pretend that all interbank exposures have the same maturity. Furthermore, I assume that the gross return on interbank deposits between date 0 and 1,  $R_{IB}$ , have the form:

$$R_{IB} = R_b - LGDk[D(1) - D(0)] \tag{20}$$

In this equation,  $R_b$  is a  $J \times 1$  vector of the required return on interbank deposits between dates 0 and 1;  $LGD$  is a diagonal matrix of the loss given default if each bank  $j$  defaults. For simplicity the loss given default for each bank is treated as a constant. Finally,  $D(1) - D(0)$  is a vector of the changes in distress experienced by each bank  $j$  between dates 0 and 1 and  $k$  is a scalar that translates the change in distress into a change in probability of default ( $\Delta PD = k[D(1) - D(0)]$ ) between

dates 0 and 1. With this specification, the return on interbank assets is equal to its required return plus an innovation that has mean 0 (changes in PD's are martingale's) and accounts for changes in expected loss on the interbank assets.

Denote bank  $j$ 's portfolio weight in interbank assets by the row vector  $\omega_{j,ib}$ , and stacking the row vectors, denote the matrix of interbank asset holdings by  $\omega_{ib}$ . Then, with this notation, the return on bank  $j$ 's assets is given by

$$R_j = \omega_{j,X}X + \omega_{j,ib}(R_b - LGDk[D(1) - D(0)])$$

If the expression for  $R_j$  is plugged into the expression for bank  $j$ 's capital at date 1 (equation (9)), and the expression for capital is plugged into the expression for bank  $j$ 's financial distress (assumption 6), it shows that each bank  $j$ 's distress when the banks are interconnected through interbank markets is a function of the distress of all of the other banks due to banks' connections through interbank deposits. Therefore, banks distress needs to be solved for as a system when they are interconnected.

To solve for bank's distress, I have linearized the system around a case in which banks are not interconnected through deposits. Tedious algebra then shows that the vector of bank's financial distress when connected through interbank markets  $D(1)_c$  is approximately equal to their distress when not connected plus adjustment terms (network multipliers) for how the network connections affect distress.

$$D(1)_c \approx (I - \Psi)^{-1}D(1)_{nc} + \text{Other Terms} \quad (21)$$

$$= D(1)_{nc} + [\Psi + \Psi^2 + \Psi^3 + \dots]D(1)_{nc} + \text{Other Terms} \quad (22)$$

Each network adjustment term depends on  $\Psi$ .<sup>43</sup> The first-order adjustment term is  $\Psi$  which captures the effect of first-order (direct) network connections. The next term captures connections between banks of length 2 in the interbank market, and so on.

The expression for  $\Psi$  is the product

$$\Psi = \Gamma\Omega_{IB}LGD, \quad (23)$$

where  $\Gamma$  is a diagonal matrix of each banks financial fragility, which captures how a change in the returns on its asset portfolio affects its distress (and also its PD) based on the date 1 realizations

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<sup>43</sup>To express  $(I - \Psi)^{-1}$  as an infinite series in  $\Psi$ , all the eigenvalues of  $\Psi$  must have a modulus less than 1.

of  $X$  when banks are not interconnected. If banks are all very well capitalized, they all have low financial fragility, the elements of the  $\Gamma$  matrix are 0, and network effects don't matter. Conversely, if some banks are not well capitalized because of bad realization of  $X$  for the bank, or just low initial capital, then network effects matter more because one bank's distress is transmitted to others through the network adjustment terms.

Recall that  $SAD$  is a weighted average of banks distress functions. When the banks have interbank deposits, then  $SAD$  is a weighted average of the elements of  $D(1)_c$ . This weighted average depends on  $X$ , capital injections into the bank at date 0 and on banks interbank deposit holdings denoted by the matrix  $\Omega_{IB}$ . The matrix of interbank deposits is typically unknown. To account for how uncertainty about it matters for choosing stress scenarios, one can proceed to identify the systemic risk factors by randomly drawing  $X$  and the elements of  $\Omega_{IB}$  from their prior distribution given the information that is publicly known. Then  $SAD$  can be computed; this process can be repeated many times, and  $SIR$  and  $COP$  can be used to identify factors that drive systemic risk, and scenarios can be generated that attempt to ensure the banking system has adequate capital given the network uncertainty that is present in the banking system.

The reason that stress scenarios may help to control the systemic risk driven by network uncertainty is because the scenarios require banks to raise more capital, and more capital reduces the effects of financial fragility, and thus reduces the effect that networks and network uncertainty has on banks distress. In particular, if some banks fragility matters more than others for network effects, then if the stress scenario emphasizes factors that affect its losses, then the capital raised in response to the scenario will help to reduce network effects. These ideas are illustrated in Figures 13 and 14. The figures illustrate the first-order and second order effects of network contagion operating through the distress channel. The main findings are when noninterbank assets produce higher returns, then interbank assets, and uncertainty about interbank exposures has a much smaller effect in assessing banks' PDs than when returns are low. Similarly, if banks are well capitalized, then network effects tend to be smaller. Conversely, if banks are poorly capitalized, then there is more uncertainty associated with network effects. The next version of this paper will attempt to formulate stress scenarios when there is uncertainty about interbank connections.

## 6 Conclusions

This paper has presented a new approach for choosing regulatory stress scenarios using dimension reduction techniques. The approach has the potential to improve regulatory stress-test practices along two dimensions. First, the paper chooses regulatory stress-scenarios to ensure that an approximate measure of systemic risk is low. Hence, the approach chooses scenarios so that the financial system is resilient to the scenarios used in the regulatory tests, as well as against a wider set of plausible scenarios. Second, the paper provides a methodology for how to incorporate information

on banks risk exposures into regulatory scenario construction. The incorporation of information and the methodology together reduce the likelihood that important directions of banks risk-taking will be missed in constructing the stress-scenario.

It is important to emphasize that the main innovations in this paper, using simulations to identify risk factors based on their ability to explain systemic risk, and then using the identified factors to develop systemic risk stress-tests are both new ideas. There is tremendous scope to further refine methods for how these ideas can be applied in practice.

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## A Proofs

**Proposition 1:** For the  $b_k$  coefficients that satisfy equation 8, each principal component  $Xb_k$  is spanned by the factors  $X\beta_k$ ,  $k = 1, \dots, K_A$ .

**Proof:** Suppose  $b$  is an eigenvector that satisfies equation (8) with positive eigenvalue  $\lambda$ .  $b$  can be written as the sum  $b^{span} + b^\perp$ , where  $b^{span}$  is from  $Xb$ 's projection onto the factors ( $Xb^{span} =$

$\sum_{k=1}^{K_A} c_k X\beta_k$ ) and  $b^\perp$  where  $Cov(Xb^\perp, X\beta_k) = b^{\perp'}\Sigma_{XX}\beta_k = 0$  for all  $k = 1, \dots, K_A$ . Equation (8) then implies that

$$\Sigma_{XX}^{-1}\Sigma_{E(X|Y)}[b^{span} + b^\perp] = \lambda[b^{span} + b^\perp] \quad (24)$$

If  $\Sigma_{E(X|Y)}b^\perp = 0$ , then

$$\Sigma_{XX}^{-1}\Sigma_{E(X|Y)}[b^{span}] = \lambda[b^{span} + b^\perp], \quad (25)$$

which implies  $b^{span} + b^\perp$  cannot be an eigenvector as in equation (8) unless  $b^\perp = 0$ . Therefore, to show solutions to (8) are spanned by the factors  $X\beta_k, k = 1, \dots, K_A$  it suffices to show:

$$\Sigma_{E(X|Y)}b^\perp = 0. \quad (26)$$

If  $[E(X|y) - E(X)]b^\perp = 0$ , then  $b^{\perp'}[E(X|y) - E(X)]'[E(X|y) - E(X)] = 0$  and  $E(b^{\perp'}[E(X|y) - E(X)]'[E(X|y) - E(X)]) = b^{\perp'}\Sigma_{E(X|y)} = \Sigma_{E(X|Y)}b^\perp = 0$ . Therefore, it suffices to show  $E(Xb^\perp|y) - E(Xb^\perp) = 0$ .

By the law of iterated expectations and the arguments of  $Y$ ,

$$E(Xb^\perp - E(Xb^\perp)|y) = E\left(E\{[Xb^\perp - E(Xb^\perp)]|X\beta_1, \dots, X\beta_{K_A}, \epsilon\}|y\right).$$

By assumption 4  $\epsilon$  is independent of  $X$ , thus it suffices to show

$$E\left(E\{[Xb^\perp - E(Xb^\perp)]|X\beta_1, \dots, X\beta_{K_A}\}|y\right) = 0.$$

It therefore is sufficient to show  $E\{[Xb^\perp - E(Xb^\perp)]|X\beta_1, \dots, X\beta_{K_A}\} = 0$ , or equivalently to show

$$E\left(E\{[Xb^\perp - E(Xb^\perp)]|X\beta_1, \dots, X\beta_{K_A}\}\right)^2 = 0$$

By assumption 5,  $E(Xb^\perp|X\beta_1, \dots, X\beta_{K_A}) = \alpha_0 + \sum_{k=1}^{K_A} \alpha_k X\beta_k$ . Therefore,

$$\begin{aligned} & E\left(E\{[Xb^\perp - E(Xb^\perp)]|X\beta_1, \dots, X\beta_{K_A}\}\right)^2 = \\ &= E[b^{\perp'}(X - E(X))'(\alpha_0 + \sum_{k=1}^{K_A} \alpha_k X\beta_k)] \\ &= b^{\perp'}\alpha_0 E(X - E(X))' + \sum_{k=1}^{K_A} \alpha_k b^{\perp'}\Sigma_{XX}\beta_k \\ &= 0. \end{aligned} \quad (27)$$

Therefore,  $b^\perp = 0$ , and  $b$  that satisfy equation (8) are spanned by  $X\beta_1, \dots, X\beta_{K_A}$ .  $\square$

**Proposition 2:** *If SAD is linearly approximated by ASAD:*

$$ASAD = C_0 + \sum_j D_j(R_j + CI_j R_f) \quad (28)$$

where  $R_j$  satisfies equation (13), and if regulators objective function takes the form

$$Prob(ASAD \geq \xi) \leq \psi$$

, then there is systemic risk factor shock  $F_A^*$  such that when the stress scenario is  $X_i = \alpha_i + F_A^* \theta_i$  for all  $X_i$ , and banks inject capital equal to the present value of their losses in the stress scenario, then after the capital is injected  $Prob(ASAD \geq \xi) \leq \psi$ .

**Proof:** *Under the conditions of the proposition*

$$\begin{aligned} ASAD &= C_0 + \sum_j D_j(R_j + CI_j R_f) \\ &= C_0 + \sum_j D_j(\alpha_j + F_A \theta_j + \epsilon_j) + \sum_j D_j CI_j R_f \\ &= C_0 + \alpha + F_A \theta + \epsilon + CIE, \end{aligned} \quad (29)$$

In the first line of equation (29), SAD is approximated as the sum of a constant  $C_0$  and linear sensitivities  $D_j$  to  $(R_j + CI_j R_f)$ . The  $D_j$  coefficients are negative since when the banks asset portfolio has higher returns, the banks net-worth increases and its distress goes down. In the last line,  $\alpha$ ,  $\epsilon$ , and CIE (capital injection equivalents) group terms involving  $\alpha_j$ ,  $\epsilon_j$ , and  $CI_j$  respectively. In the above expression,  $F_A \theta + \epsilon$  is a single random variable. Let  $H(\cdot)$  denote its CDF.

The rest of the proof has two steps. The first step solves for the minimum CIE, denoted  $CIE^*$  such that if banks inject enough capital to achieve  $CIE^*$ , then  $Prob(SAD \geq \xi) \leq \psi$ . The second step solves for a stress scenario that requires banks to inject enough capital to achieve  $CIE^*$ .

*Step 1: Solve for  $CIE^*$ .*

*Manipulation of equation (29) shows*

$$Prob(SAD \geq \xi) = 1 - H(\xi - C_0 - \alpha - CIE). \quad (30)$$

*The  $CIE^*$  that solves*

$$1 - H(\xi - C_0 - \alpha - CIE^*) = \psi,$$

*will be the smallest CIE that satisfies the condition  $Prob(SAD \geq \xi) \leq \psi$ . Manipulating the above*

equation, it is given by:

$$CIE^* = C_0 + \alpha - \xi - H^{-1}(1 - \psi). \quad (31)$$

Note that since  $D_j(\cdot) < 0$  in equation 29, if more capital needs to be injected into the banking system, then  $CIE^* < 0$ .

*Step 2: Solve for  $F_A$  and the stress-scenario.*

If the systemic risk factors are set to value  $F_A^*$ , then in the resulting stress-scenario, each banks gross return is  $R_j^* = \alpha_j + F_A^* \theta_j$ . The amount of capital banks need to raise in the scenario is equal to their losses  $-A_j(R_j^* - 1)$  less any excess capital that was previously held. The excess capital just adds constant terms to the analysis. For simplicity, excess capital is assumed to be zero. As a fraction of its initial assets  $A_j$ , each bank requires additional date 1 capital  $\Delta Cap_j = -(R_j^* - 1)$ . Substituting the date 1 capital raised as a fraction of assets into equation (29) for  $CI_j R_f$  shows the capital raised alters SAD at date 1 by the amount

$$\begin{aligned} \Delta SAD &= -\sum_j D_j (R_j^* - 1) \\ &= -(\alpha + F_A^* \theta - \sum_j D_j) \end{aligned}$$

If  $F_A^*$  is chosen so that  $\Delta SAD = CIE^*$ , then the resulting stress scenario will require that banks inject enough capital for date 1 so that SAD at date 1 is reduced by the amount  $CIE^*$ . Therefore,  $F_A^*$  must be chosen so that

$$F_A^* \theta = -CIE^* - \alpha + \sum_j D_j. \quad (32)$$

If there is only a single systemic risk factor, then  $\theta$  is a scalar, and

$$F_A^* = \frac{-CIE^* - \alpha + \sum_j D_j}{\theta}.$$

If there are several systemic risk factors, then there are an infinite number of solutions for  $F_A^*$ , all of which satisfy equation (32).

To verify that this is the correct solution, note that as a result of the stress scenario, at date 0 each bank  $j$  will be required to inject capital  $CI_j = \frac{-(\alpha_j + F_A^* \theta_j - 1)}{R_f}$  as a fraction of its date 0 assets. Since the capital is invested in risk free assets, it will grow to  $-(\alpha_j + F_A^* \theta_j - 1)$  at date 1. Plugging

into  $SAD$  and summing across  $j$ ,  $SAD$  is changed by the amount

$$\begin{aligned}
\Delta SAD &= -\sum D_j(\alpha_j + F_A^*\theta_j - 1) \\
&= -(\alpha + F_A^*\theta - \sum_j D_j) \\
&= -\left[\alpha + \left(\frac{-CIE^* - \alpha + \sum_j D_j}{\theta}\right)\theta - \sum_j D_j\right] \\
&= CIE^*
\end{aligned}$$

## B Sliced Inverse Regression (SIR) and Correlation Pursuit (COP)

The purpose of the following two subsections is to provide additional information on the statistical interpretation of SIR, and details on the implementation of SIR and COP. A full description of the SIR methodology is described in Li (1991), Chen and Li (1998), and COP is described in Zhong et al (2012).

The  $R$  package *dr* contains a module for computing sliced inverse regression and related methods [Weisberg (2014)].<sup>44</sup> An  $R$  package for COP can be downloaded from “<http://cran.r-project.org/web/packages/COP>” [Zhong et al (2012)].

### B.1 Sliced Inverse Regression

Section A shows that the principal components identified by SIR span a subspace of the space spanned by the systemic risk factors. The purpose of this subsection is to illustrate the relationship between the uncovered principal components and systemic risk. The exposition closely follows Chen and Li (1998). They show SIR can be interpreted as solving for a first principal direction  $b_1$ , that maximize the squared correlation between  $X'b_1$  and a possibly nonlinear transformation  $T(Y)$  of the  $Y$  variables:

$$\max_{T(Y), b_1} Corr(T(Y), X'b_1)^2.$$

For given  $b_1$ , the transformation of  $Y$  that is most correlated with  $X'b_1$  is  $E(X'b_1|Y)$ . The squared correlation is given by

$$\left(\frac{Cov\{[E(X'b_1)|Y], Xb_1\}}{\sqrt{b_1'\Sigma_{XX}b_1 \times Var[E(X'b_1|y)]}}\right)^2 = \frac{b_1'\Sigma_{E(X|Y)}b_1}{b_1'\Sigma_{XX}b_1}$$

<sup>44</sup>See <http://cran.r-project.org/web/packages/dr/vignettes/overview.pdf>

The correlation is homogeneous of degree 0 in  $b_1$ . Therefore, restricting  $b_1$  so that  $b_1' \Sigma_{XX} b_1 = 1$  is without loss of generality. Solving for optimal  $b_1$  subject to the restriction then implies  $\Sigma_{E(X|Y)} b_1 = \lambda_1 \Sigma_{XX} b_1$ . Substituting in for  $\Sigma_{E(X|Y)} b_1$ , the squared correlation can be written as

$$\text{Corr}(E(Xb_1|Y), Xb_1)^2 = \frac{b_1' \Sigma_{XX} b_1 \lambda_1}{b_1' \Sigma_{XX} b_1} = \lambda_1.$$

Therefore  $Xb_1$  has maximal squared correlation  $\lambda_1$  with a transformation of systemic impairment. Each additional principal direction  $b_k$  is orthogonal to the preceding principal directions, and maximizes  $\text{Corr}^2(T(Y), X'b_k)$ .

To perform SIR, it is necessary to compute  $E(X|Y)$ . As noted by Li (1991), this could be done by nonparametric regression, but to ease the computational burden, it is simply done by binning the data and taking sample averages within bins. More specifically, let  $Z$  be an  $N \times P + 1$  partitioned matrix of the data ( $Z = [Y, X]$  where  $Y$  is  $N \times 1$  and  $X$  is  $N \times P$ ) that has been sorted by  $Y$  and partitioned into  $M$  bins that each have  $S$  rows, with  $Z_m = (Y_m, X_m)$ . The function  $E(X|Y)$  is approximated by the sample average of the  $X$ 's in each bin:

$$\bar{X}_m = \frac{1}{S} \sum_{s=1}^S X_m(s, \cdot).$$

In addition, the unconditional expected value of  $X$  is estimated by  $\bar{X} = \frac{1}{M} \bar{X}_m$ , and  $\text{Var}(E(X|Y))$  by

$$\hat{\Sigma}_{E(X|Y)} = \frac{1}{M} \sum (\bar{X}_m - \bar{X})' (\bar{X}_m - \bar{X}),$$

and  $\text{Var}(X)$  is estimated by

$$\hat{\Sigma}_{XX} = \frac{1}{N} \sum_{i=1}^N (X_i - \bar{X})' (X_i - \bar{X}).$$

The principal component factors are found by plugging the sample estimates into equation (8) and then solving for the eigenvalues and  $b_k$  coefficients.

Li (1991) shows that SIR's ability to identify the space spanned by the factors is robust to the size of the bin-size  $S$ . However, to choose the number of variables to use in SIR, Zhong et al (2012) finds the number of slices matter for performance. They find that choosing  $S = 20$  performs well. For this reason in the analysis in this paper chooses  $S = 20$  observations per slice.

## B.2 Correlation Pursuit

Correlation pursuit chooses the variables to include in SIR to create factors. The universe of variables considered is denoted  $X$  where  $X$  is an  $N \times P$  matrix of  $N$  realizations of  $P$  random variables. COP assumes that the number of factors is  $K$ . How  $K$  is chosen will be discussed below.

COP begins with a randomly chosen set of  $K + 1$  variables  $A \in X$ . The variables that are not in  $A$  are denoted  $A^C$ . To find the variables that are most suitable for creating factors to explain  $T(Y)$ , COP then scrolls through all  $P$  variables, and in doing so performs either a variable addition step, in which a variable is added to the set  $A$ , or a variable deletion step in which a variable is taken away. For a given  $K$ , the procedure continues until no more variables can be added or deleted. The final set of variables  $A(K)$  is the set of variables that is chosen under the assumption that there are  $K$  factors.

**Variable addition step:** To perform the variable addition step, let  $t$  denote a candidate  $X$  variable being considered for addition to  $A$ . To determine if the variable should be added,  $COP$  creates a test based on the scaled improvement in each of the  $K$  eigenvalues associated with the principal components when variable  $t$  is added.

The scaled improvement in the  $i$ 'th eigenvalue is denoted  $COP_i^{A+t}$  given by

$$COP_i^{A+t} = N \frac{(\lambda_i^{A+t} - \lambda_i^A)}{1 - \lambda_i^A}, \quad (33)$$

where the superscripts  $A$  and  $A + t$  denote the sets of variables used in computing the eigenvalues. Because the  $\lambda_i$  coefficients have the interpretation the  $R^2$  from using the  $i$ 'th factor to explain  $T(Y)$ , the  $COP_i^{A+t}$  statistics resemble  $F$ -tests for whether the addition of the variable  $t$  statistically improves the predictability attributable to a factor. The sum of these statistics is denoted  $COP_{1:K}^{A+t}$  ( $= \sum_{i=1}^K COP_i^{A+t}$ ). The statistic

$$\overline{COP}_{1:K} = \max_{t \in A^C} COP_{1:K}^{A+t};$$

and  $X_{\bar{t}}$  is a variable that attains this maximum if added to  $A$ . The variable  $X_{\bar{t}}$  is added to  $A$  if  $\overline{COP}_{1:K} > c_e$ , where  $c_e$  is a critical value for determining whether a variable should be added.

**Variable deletion step:** The variable deletion step is analogous to the the addition step. The scaled deterioration in the explanatory power of factor  $i$  from deleting the variable  $t$  from the set  $A$  is given by

$$COP_i^{A-t} = N \frac{(\lambda_i^A - \lambda_i^{A-t})}{1 - \lambda_i^A}. \quad (34)$$

The statistic  $COP_{1:K}^{A-t} = \sum_{i=1}^K COP_i^{A-t}$  measures the deterioration in fit from deleting the variable  $t$ . The statistic  $\underline{COP}_{1:K}$  denotes:

$$\underline{COP}_{1:K} = \min_{t \in A} COP_{1:K}^{A-t}, \quad (35)$$

and variable  $X_{\underline{t}}$  attains this minimum. If  $COP_{1:K} < c_d$ , then variable  $X_{\underline{t}}$  is deleted from  $A$ .

Zhong et al (2012) derive conditions under which as  $N$  goes to infinity with slice size ( $=$  bin



size) fixed, COP consistently chooses  $X$  variables that should be used for SIR and consistently deletes variables that should not be used. When using COP in finite samples, they recommend choosing the critical values  $c_e$  and  $c_d$  through five-fold cross-validation.

The above shows how to select the variables for a given number of factors  $K$ . To choose  $K$ , Zhong et al (2012) used a BIC type information criterion based on Zhu et al (2006).

## C SCSA Applied to Rates Positions

In Section 4 the SCSA approach is applied to six banks that hold portfolios of zero coupon bonds in 8 countries. The country list and the specific zero coupon bonds are listed in the text.

### C.1 Creating banks bond holdings

Banks' holdings of zero coupon bonds are generated randomly drawing a DVO1 for each bond position, and then inverting the DV01 to compute the amount of money invested in the bond. The DV01 values are drawn from two different distribution functions. In our first application, each DV01 is distributed  $N(-.5, 1)$ , where the DV01 approximates the change in the dollar value of an exposure due to a one-basis point increase in yield. After the dollar amount of money invested in each bond is computed, the amounts invested in each bond are converted to portfolio weights that sum to 1. For the second choice of distribution functions, the exposures are randomly simulated by choosing DV01s for each yield from a  $N(0, 1)$  distribution.

To illustrate the approach, we focus on simulating a position in a single zero coupon bond. The approach for each bond is analogous. Let  $y(t)$ , and  $P(t)$  denote the price and yield of a continuously compounded zero coupon bond with maturity  $t$ . Let  $n(t)$  be the number of zero coupon bonds of maturity  $t$  in the portfolio. The value of the holdings of zero coupon bonds is

$$v(t) = n(t)P(t) = n(t)e^{-y(t)t}$$

The DV01 for the position is measured as the first-order approximation of the change in its value due to a 1 basis point increase in its yield:

$$\begin{aligned} DV01(t) &= n(t) \frac{\partial P(t)}{\partial y(t)} \times .0001 \\ &= -n(t)tP(t) \times .0001 \\ &= -tv(t) \times .0001 \end{aligned}$$

Recall the  $DV01(t)$  is randomly simulated. Rearrangement of the last line shows how to convert the simulated  $DV01(t)$  to  $v(t)$ :

$$v(t) = \frac{10^4 DV01(t)}{t}$$

The same approach is used to simulate the investments in all zero coupon bonds in countries indexed by  $c$  and for maturities indexed by  $t$ , with corresponding values  $v_c(t)$ ,  $c = 1, \dots, C$ ,  $t = 1, \dots, T$ .

For the purposes of the further analysis, the investments are expressed as portfolio weights, i.e. the portfolio weight for bank  $i$  in country  $c$  with maturity  $t$  is  $\omega_c^i(t) = \frac{v_c(t)}{\sum_{c,t} v_c(t)}$ . The vector of portfolio weights for bank  $i$  is its stacked vector of weights in each country: denoted  $\omega^i$ .<sup>45</sup>

## C.2 Data and Simulating X Variables

The  $X$  variables in the analysis are changes in the zero coupon bond yields and/or returns of Australia, Canada, Germany, Japan, Sweden, Switzerland, Great Britain, and the United States. The tenors used in the analysis ranged from short maturity to 15 years for all countries, but was as long as 30 years for the US.<sup>46</sup> The yield curve data on these variables spans the period from February 2006 to October 2013.<sup>47</sup>

Unless otherwise indicated, the time horizon for the analysis is one-month and yield curve changes are simulated over a one-month horizon using historical simulation. With this methodology, it is assumed the distribution of future shocks to the yield curve are the unconditional distribution of historical yield curve changes in the data; i.e. the shocks are the time series of 165 non-overlapping monthly changes in 83 zero-coupon yields that occurred over the time span of the data.

## C.3 Specifying the SAD objective function

Recall that SAD is the weighted average of banks distress functions. In equation (10), the functional form of banks distress functions is left unspecified. In the empirical analysis we used distress functions that have a logit form for each bank  $j$ :

<sup>45</sup>Because 83 zero coupon bonds are used in the analysis,  $\omega^i$  is  $83 \times 1$ .

<sup>46</sup>For the United States the yield tenors in years are 1,2,3,4,5,7,8,10,12,15,20,25,30. For the other countries the tenors in years are 0.25,.5,.75,1,2,3,5,7,10,15.

<sup>47</sup>Data on zero coupon yields for all countries other than the US was generously provided by the International Finance Division at the Federal Reserve Board. Zero coupon yield data for the U.S. is based on the methodology from Gurkaynak et. al.(2006). This data is frequently updated and is available along with their working paper on the Federal Reserve Board's website.

$$D_j(\cdot) = \frac{1}{1 + \text{Exp}(a_j + b_j C_j(1)/\sigma(C_j(1)))}, \quad (36)$$

where  $C_j(1)$  is bank  $j$ 's capital ratio at date 1 and  $\sigma(C_j(1))$  is the standard deviation of the capital ratio as of date 1. The expression  $C_j(1)/\sigma(C_j(1))$  is bank capital divided by a measure of risk of the capital, and hence can be interpreted as risk weighted capital since for a given capital ration, the greater is the risk the lower is the effective risk-weighted capital and the greater is a banks financial distress.<sup>48</sup>

The parameters  $a_j$  and  $b_j$  should ideally be calibrated so that the distress function for each bank roughly captures the relationship between the banks intermediation capacity and its risk-weighted capital. For the purposes of this paper the parameters are not calibrated; instead they are set so that  $a_j = 0$  and  $b_j = .95$  for all  $j$ . Because capital ratios span from 0 to 1, with this specification, distress for each bank has a lower bound of 0, which is approached as the volatility of capital goes to 0. The upper bound for distress is 0.5 which is approached as capital goes to 0. This means an insolvent bank loses half of its maximal financial intermediation capacity. the other parameter that needed to be calibrated as part of the analysis was the choice of banks initial capital ratios. For the analysis here this ratio was chosen so that banks have a two-percent chance of bankruptcy over the stress horizon. This is not a realistic choice, but was made for convenience and should be altered in the future.<sup>49</sup>

## D Simulation of Interest Rate Changes Based on Wright (2011).

The estimation of a multi-country term structure models is well beyond the scope of this paper. To simulate multi-country term structures, I use parameter estimates that were generously provided by Jonathan Wright (Wright (2011)). Wright (2011) estimates a homoskedastic discrete-time affine term structure models for Australia, Canada, Germany, Japan, New Zealand, Norway, Sweden, Switzerland, the United Kingdom, and the United States.

In Wright's model, in each country  $c$ ,  $n$ -month zero-coupon bond prices at date  $t$  are an affine function of a vector of country specific state variables  $X_{c,t}$  at date  $t$ .

$$P_{c,t}^n = \exp(-A_{n,c} - B'_{n,c} X_{c,t}) \quad (37)$$

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<sup>48</sup>An alternative approach could use Basel risk-weights. An advantage of the approach in this paper is that Basel risk-weights for simplicity assume risk scales linearly, but in fact this is not descriptive of how risk scales.

<sup>49</sup>This seems like a high but defensible choice of unconditional insolvency risk over a year, but not over a month. It is set to be high in the analysis so that the choice of capital can be calibrated on the basis of the historical bond returns during our short sample of 165 months. A more realistic default probability per month is perhaps .0017, or a bit less than two in a thousand. A sample of 165 months cannot be used to calibrate the amount of capital needed to cover losses with such a low probability. Instead a parametric model of returns in the tail would need to be fit to the data.

There are 5 state variables for  $X_c$  in each country. The first three are the the first three-principal components of the zero coupon bond yield curve in each that country for maturities from 3 months to 10 years. The fourth and fifth state variables are exponential weighted moving averages of inflation and GDP growth in each country.

The physical and dynamics of the state variables in each country are modeled as following a VAR(1):

$$X_{c,t+1} = \mu_c + \Phi_c X_{c,t} + \Sigma_c \epsilon_{c,t+1} \quad (38)$$

where  $\epsilon_{c,t+1}$  are a vector of Gaussian white noise innovations for each country, and  $\Sigma_c$  is a lower Cholesky matrix.

The dynamics of state-variables under the pricing measure follow a VAR(1) but with a different set of “starred” parameters:

$$X_{c,t+1} = \mu_c^* + \Phi_c^* X_{c,t} + \Sigma_c \epsilon_{c,t+1} \quad (39)$$

The model imposes the restriction that GDP growth and inflation are unspanned term structure factors. This means that contemporaneous realizations of those variables do not affect the realizations of the contemporaneous term structure of interest rates. But, the evolution of those variables does affect the future realizations of the principal components that do affect the term structure. Further details on the interest rate model, its foundation, restrictions, and its estimation are in Wright (2011).

Wright provided me with his parameter estimates and data. For each country  $c$  using the estimates  $\mu_c$ ,  $\Phi_c$ ,  $\Sigma_c$ , and data  $X_{c,t}$ ,  $t = 1, \dots, T$ , I solved equation 38 for the realizations of  $\epsilon_{c,t}$ ,  $t = 1, \dots, T$ .

In this paper I modified the model in Wright slightly for the purposes of simulating future term-structures across countries. Specifically, Wright estimated the model for each country separately without imposing restrictions on how term structures are related across countries. For the purposes of this paper, I wanted the model to incorporate at least in a rough way how the term-structures across countries are related. To do so, I assumed the  $\epsilon_c$  vectors are contemporaneously correlated across countries. Specifically, let  $\epsilon_{C,t} = (\epsilon'_{c_1,t}, \dots, \epsilon'_{c_{10},t})'$  denote the stacked vector of countries'  $\epsilon_c$  realizations at date  $t$ .

I assume

$$\epsilon_C \sim \text{i.i.d.} N(0, \Omega_C) \quad (40)$$

where  $\Omega_C$  is a  $50 \times 50$  block diagonal matrix.

Because the data-set is not a balanced panel (see Wright (2011)), each  $c_i, c_j$  block of the covariance matrix was estimated as

$$\hat{\Omega}_{c_i, c_j} = \frac{1}{T_{i,j}} \sum_{t=1}^{T_{i,j}} \epsilon_{c_i, t} \epsilon'_{c_j, t},$$

where  $T_{i,j}$  are the number of time-series observations that overlap for countries  $i$  and  $j$ .

The resulting matrix  $\hat{\Omega}$  was not positive definite because it had a few very tiny but negative eigenvalues. To impose positive definiteness on the estimated  $\Omega_C$  matrix, the estimate  $\tilde{\Omega}_C$  was used as the estimate of  $\Omega_C$ :

$$\tilde{\Omega}_C = v_3 D_3 v_3' + \tilde{D}, \quad (41)$$

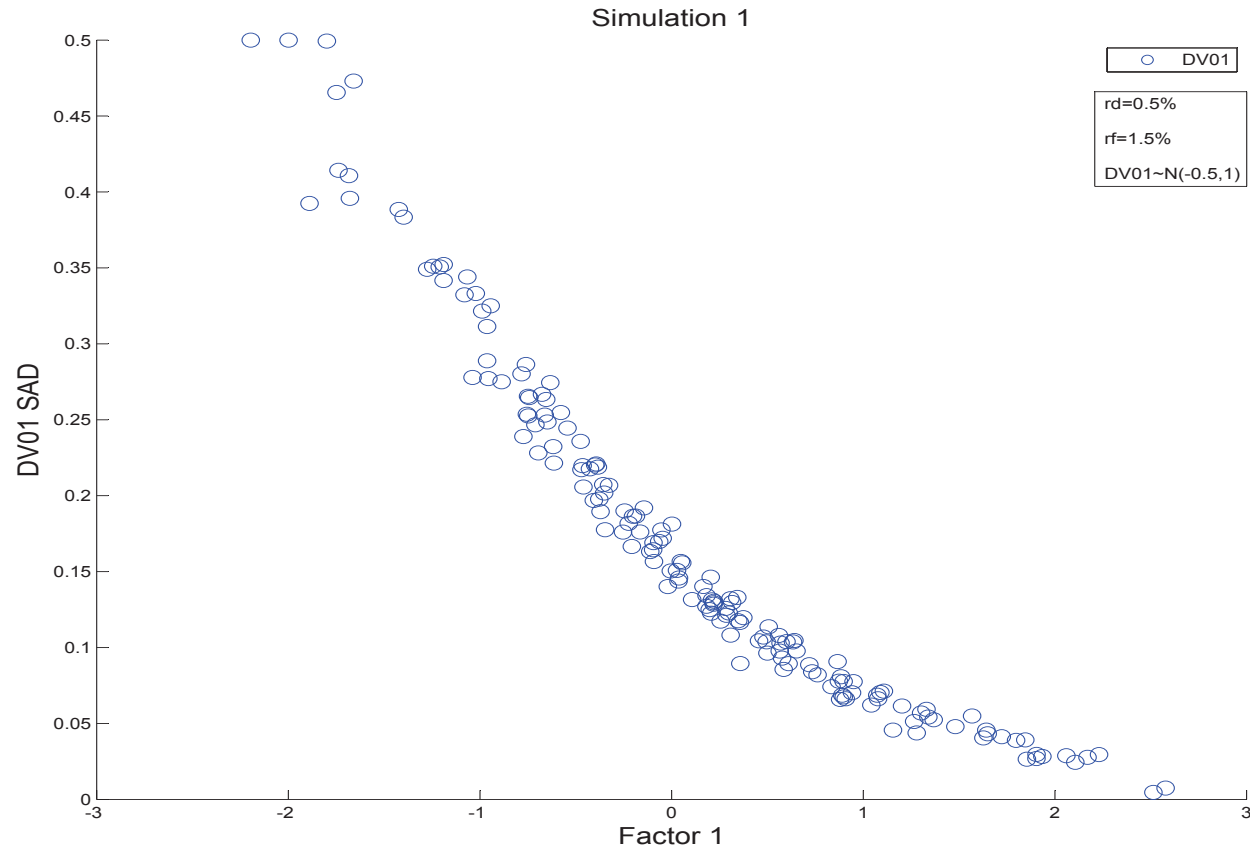
where  $v_3$  is a  $50 \times 3$  matrix of the first three eigenvectors of  $\hat{\Omega}_C$ ,  $D_3$  is a diagonal matrix of the first three eigenvalues, and  $\tilde{D}$  is a  $50 \times 50$  diagonal matrix whose parameters were chosen so that the diagonal elements of  $\tilde{\Omega}_C$  match the diagonal elements of  $\hat{\Omega}_C$ .

To simulate yield curves using this framework, beginning with initial values of the state variables in each country at date  $t$ ,  $\epsilon_{C,t+1}$ ,  $\epsilon_{C,t+2}$  and so on are drawn from their estimated distribution, and then the state variables simulated forward using their physical dynamics from equation (38). Then, zero-coupon bond yields and prices in each country are computed using the formula (37).<sup>50</sup>

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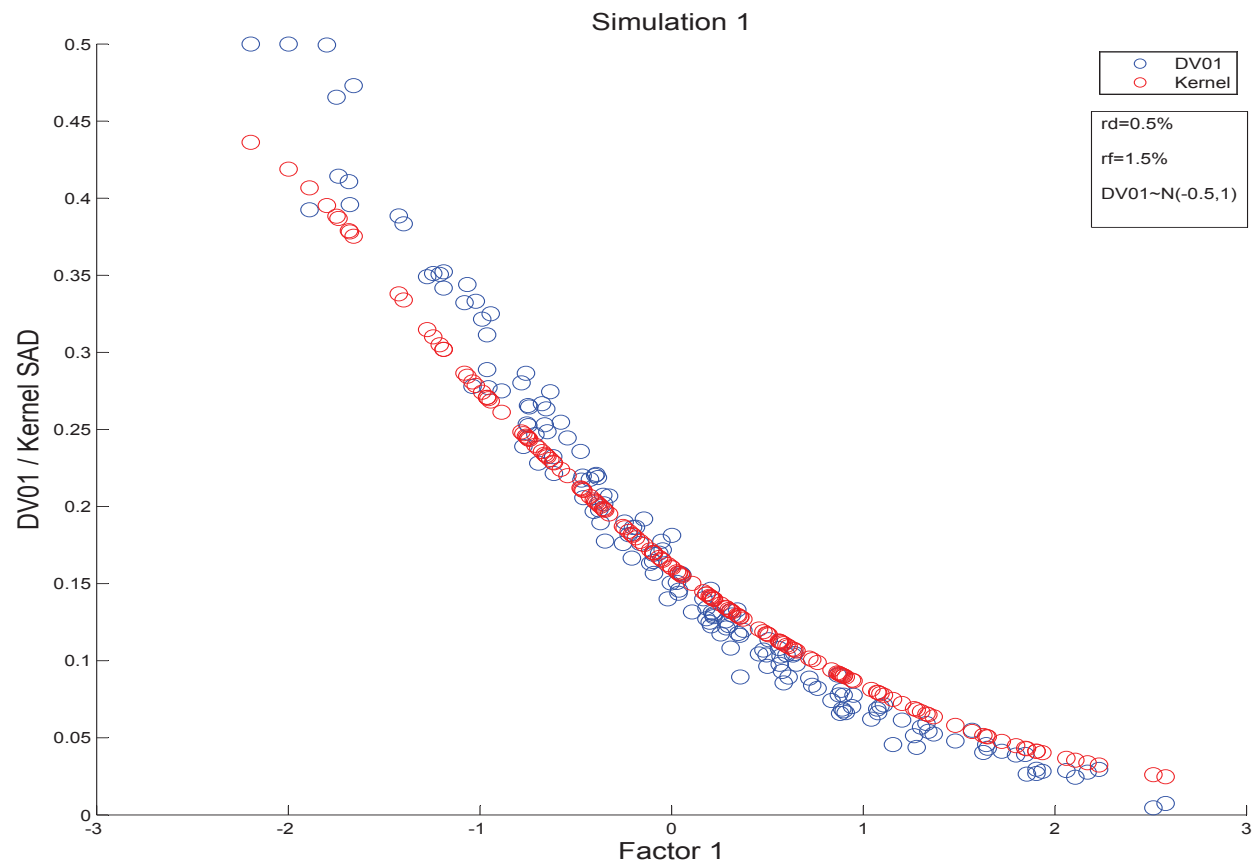
<sup>50</sup>The  $A_n$  and  $B_n$  parameters are the solutions of Riccati difference equations that are solutions of equations based on parameters for the evolution of the state variables under the  $Q$  measure. For details, see Wright (2011).

Figure 1: Scatter Plot of SAD versus extracted Systemic, Factor Long on Average Portfolio



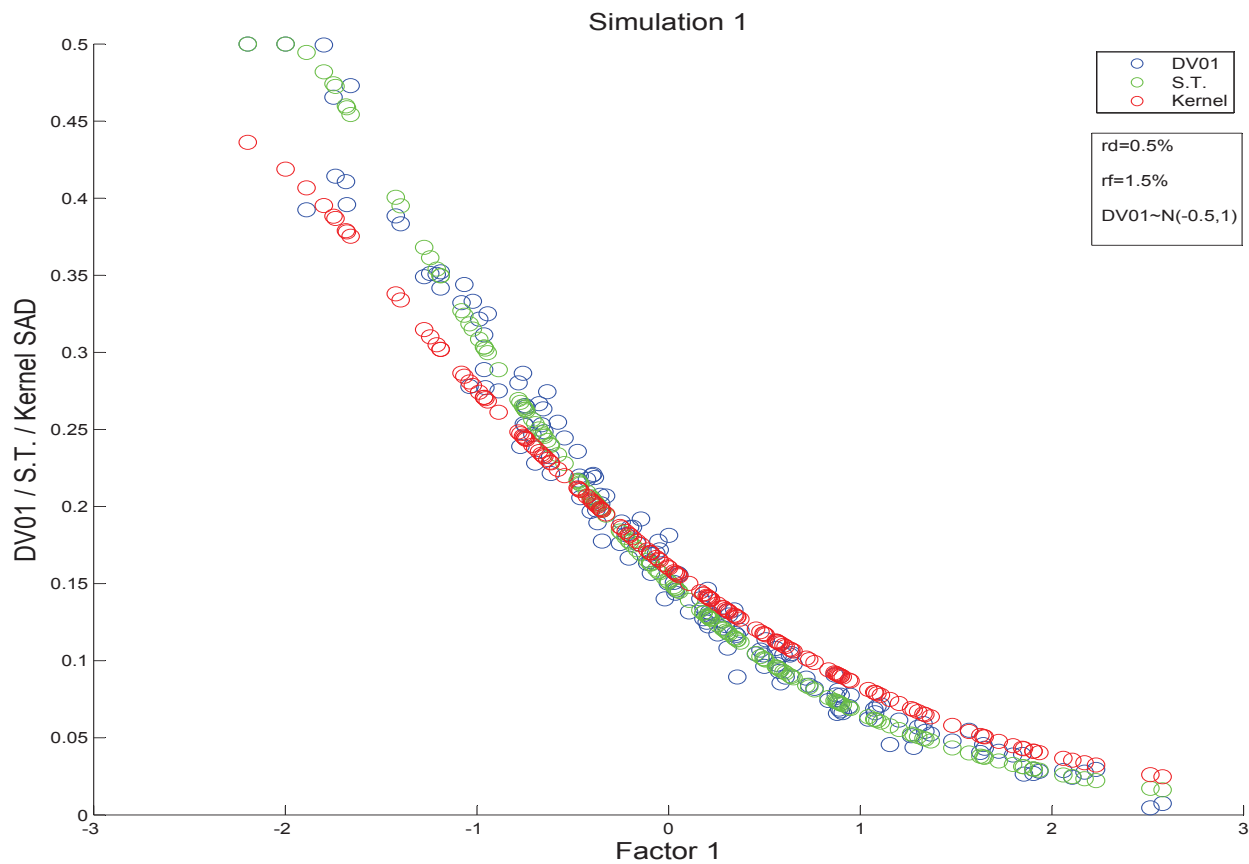
*Notes:* For a randomly chosen set of interest-rate portfolios for 6 banks (simulation 1), and for a large number of changes in zero coupon yield curves in 8 countries, the figure presents a scatter plot of a measure of System Assets in Distress (DV01 SAD), a measure of the banks joint financial distress, against a systemic risk factor (Factor 1) that was extracted to explain banks distress. The figure shows the extracted factor tracks SAD well. SAD in the figure is approximated using linear sensitivities to interest rate changes, measured by DV01's at each point on the yield curves of all 8 countries. The portfolios have long interest-rate exposure (positive duration) on average; the time horizon over which portfolio returns is computed is 1-month. The one month annualized risk-free rate,  $r_f$  and rate on deposits,  $r_d$  are described as in the figure; as are DV01 risk-sensitivities. Additional details are in section 4 of the text.

Figure 2: Nonparametric Regression of SAD versus extracted Systemic Factor, Long on Average Portfolio



*Notes:* For the randomly chosen portfolios from Figure 1, the figure presents the same scatter plot as in Figure 1, and overlays the scatter plot with a nonparametric kernel regression (red dots) of the relationship between the systemic risk factor and SAD. The regression further confirms that the factor tracks SAD well and shows the relationship between SAD and the factor is slightly nonlinear. Additional details are in Figure 1 and section 4 of the text.

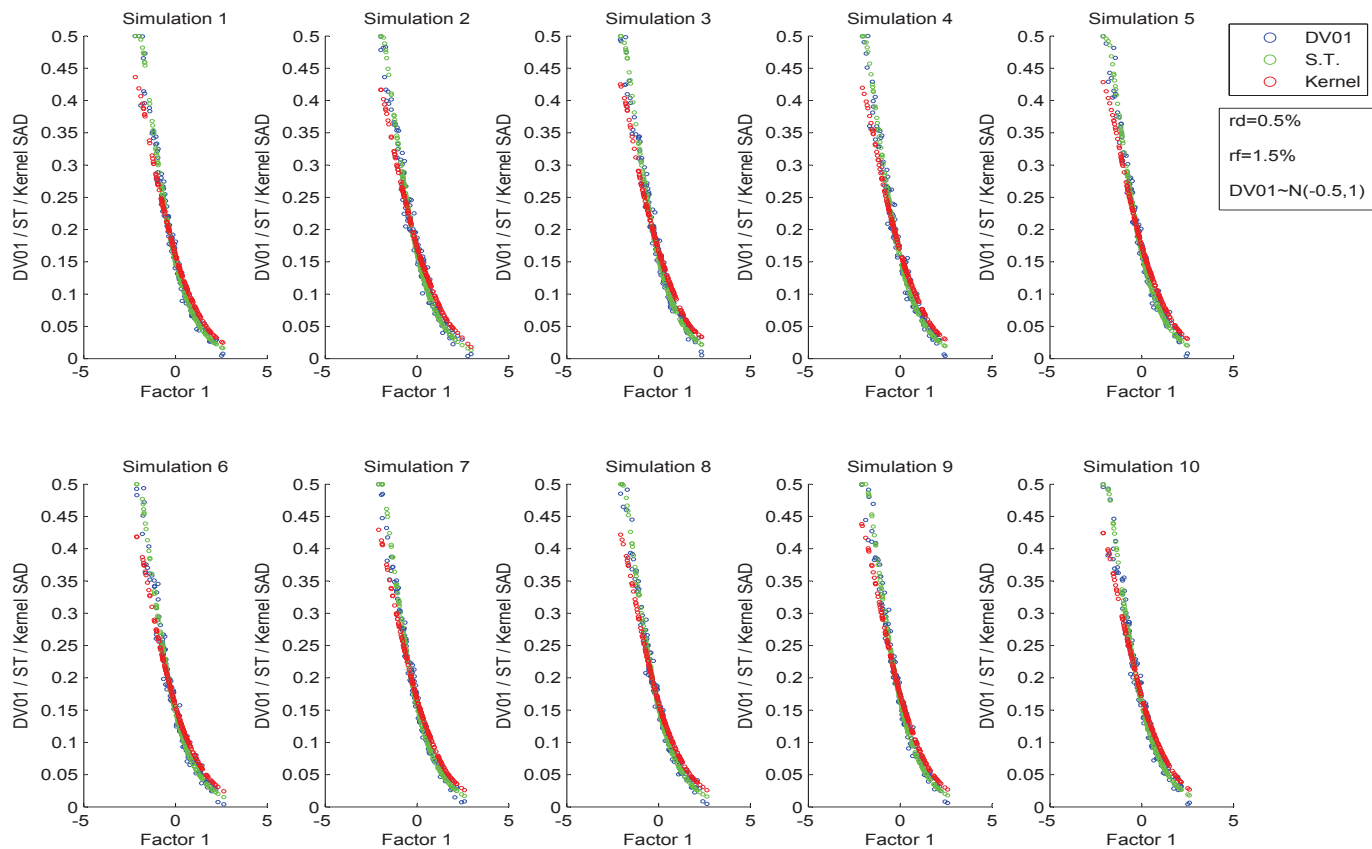
Figure 3: Stress Tests Based on Systemic Factor: Relation to SAD, Long on Average Portfolio



*Notes:* For the randomly chosen portfolios from Figures 1 and 2, the figure presents the same scatter plot as in Figure 2, and overlays the scatter plot with a measure of what *SAD* would be (the green dots labeled S.T. in the figure), if a stress test (S.T.) was created for each realization of Factor 1. The figure shows that *SAD* based on the stress-tests tracks true *SAD* well. This suggests that ensuring banks are well capitalized against a stress-test based on the factor has the potential to ensure the banks in the figure can be well protected against increases in *SAD*. Additional details are in Figures 1 and 2 and section 4 of the text.

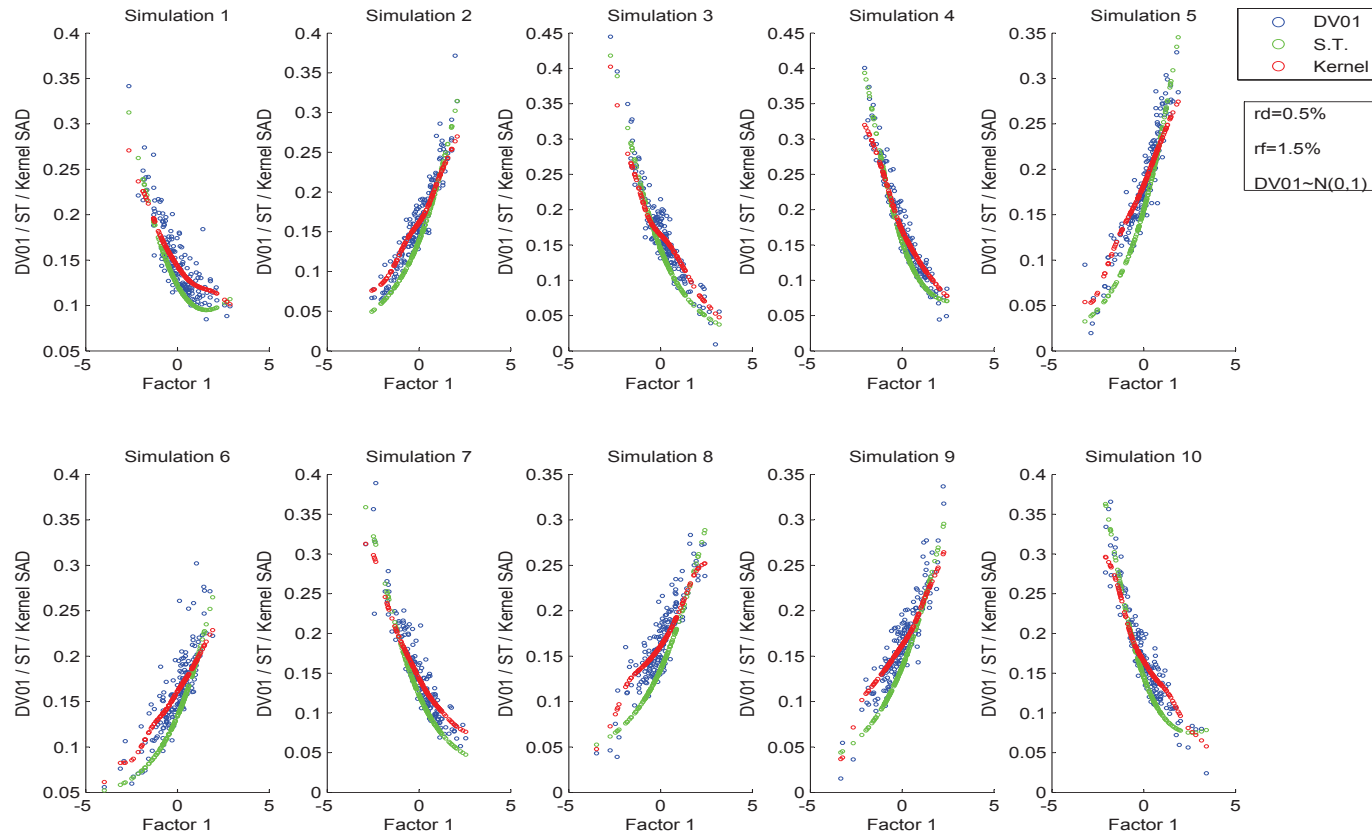


Figure 4: Stress Tests Based on Systemic Factor: Relation to SAD, Simulated Portfolio Sets 1 - 10, Long on Average Portfolio



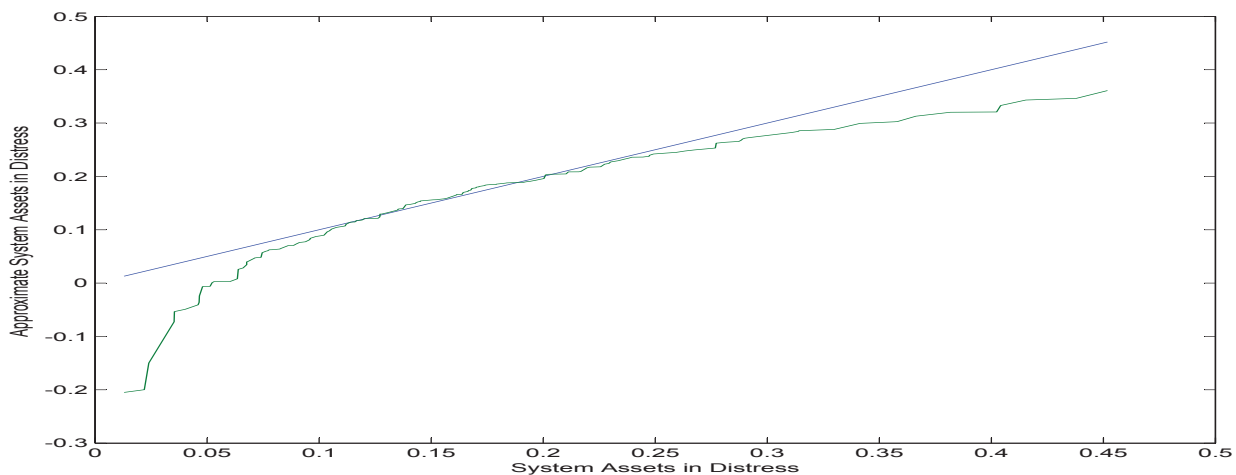
Notes: For ten simulations (simulation 1 - simulation 10) of randomly chosen portfolios for 6 banks, the figure presents information that is comparable to Figure 3. All ten simulated portfolios are created using the same approach as was used in Figure 1. The figure shows that the results from the ten sets of simulated portfolios produce results that are similar to the results in Figure 1 to Figure 3.

Figure 5: Stress Tests Based on Systemic Factor: Relation to SAD: Simulated Portfolio Sets 1 - 10, Neither Long Nor Short Portfolio



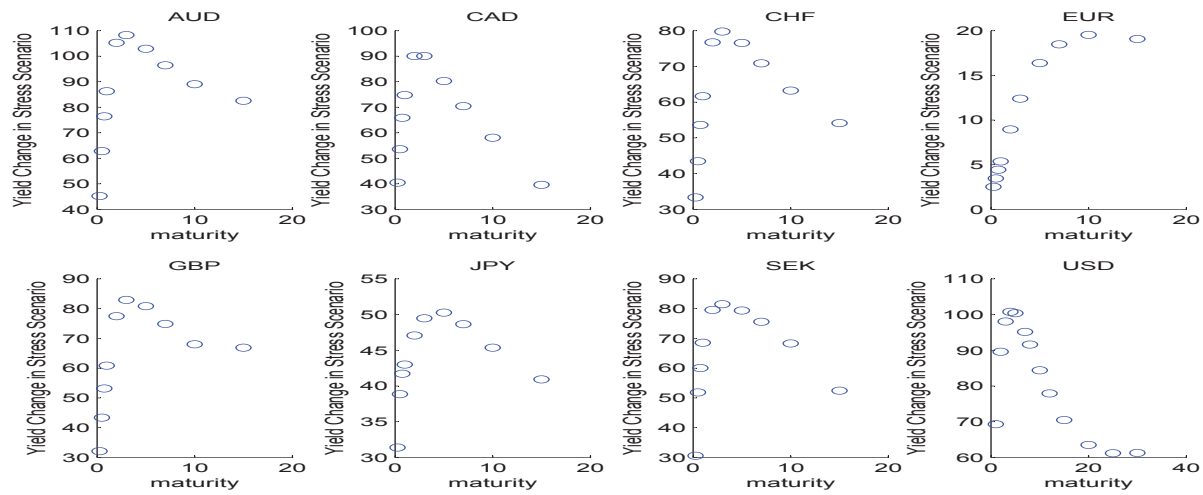
*Notes:* For ten simulations (simulation 1 - simulation 10) of randomly chosen portfolios for 6 banks, the figure presents information that is comparable to Figure 3. All ten simulated portfolios are created using the same approach as was used in Figure 1 except that the ten portfolios are neither long nor short interest rates on average. The figure shows that the results from the ten sets of simulated portfolios produce results that are qualitatively similar to the results in Figure 4, but the ability of the factors to track SAD, and the ability of stress-tests based on the factors to track SAD is not quite as strong because the directionality of the interest rate exposure is not quite as uniform as in the long-biased portfolios. Details on how these portfolios are constructed are contained in Figures 1 - 3, and section 4.

Figure 6: QQ-Plot of Approximate SAD vs SAD: Simulated Portfolio 1



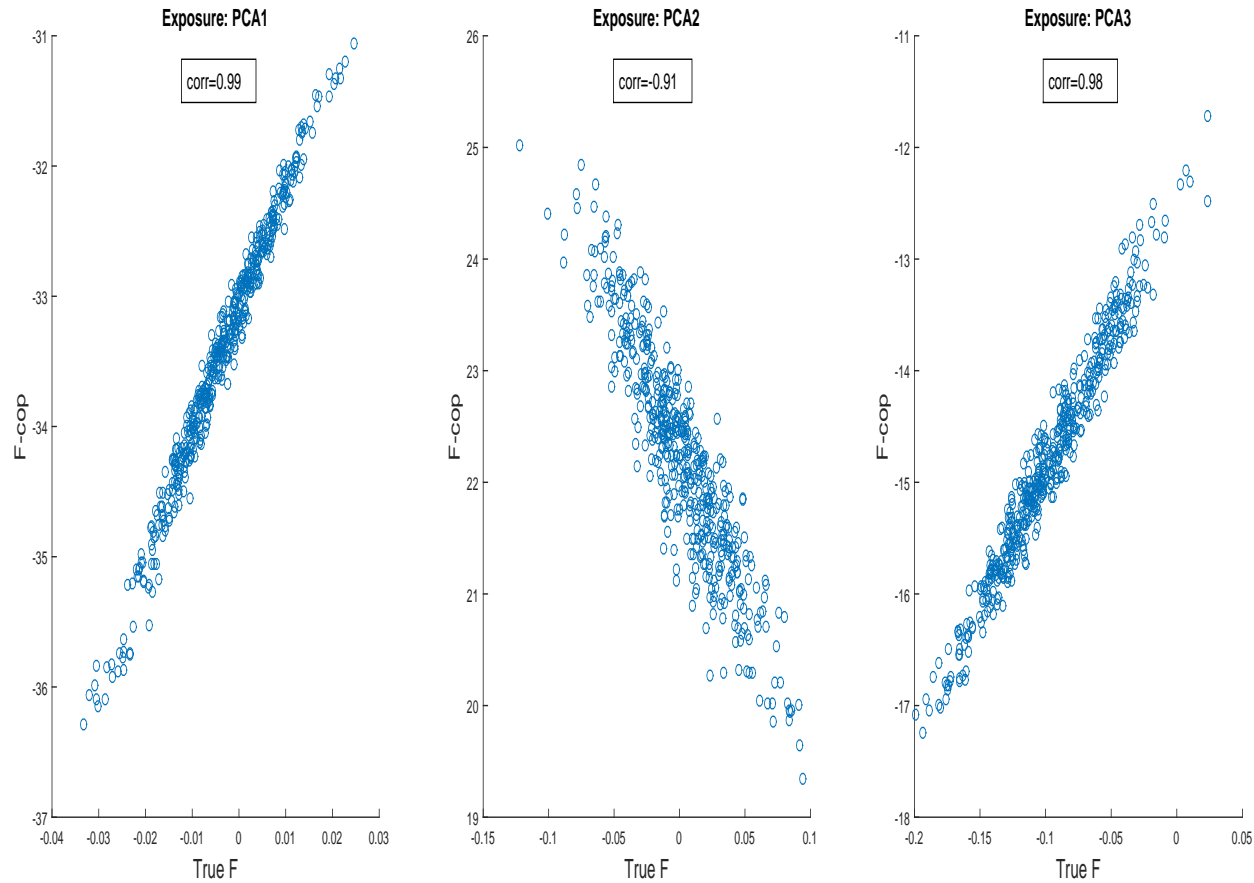
*Notes:* For the simulated portfolios for 6 banks in Figure 1, the Figure presents a QQ-plot of SAD vs Approximate SAD, where Approximate SAD is computed by linearizing SAD. The QQ plots shows that while SAD and the approximation are correlated, the distribution of SAD and the approximation of SAD differ in the tails. As discussed in the text, this means that in practice that approximate SAD helps to find the directions and magnitude to stress-variables in a stress-scenario, but to achieve a SAD objective the magnitude of the stresses needs to be adjusted based on SAD.

Figure 7: Optimal Stress-Scenario: Simulated Portfolio 1



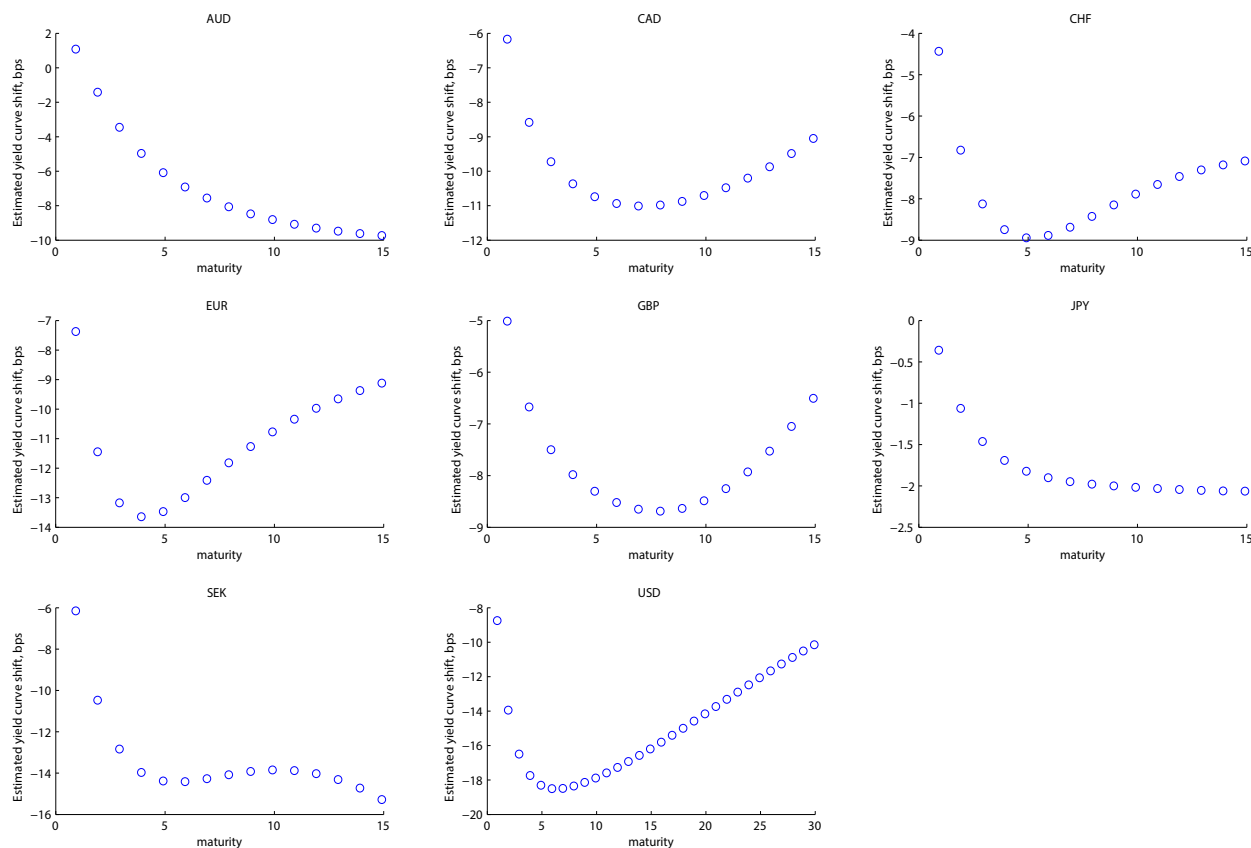
Notes: For the simulated portfolios for 6 banks in Figure 1, the Figure presents an interest-rate risk stress scenario. The scenario instantaneously shifts the default-free yield curves in the 8 countries considered by the number of basis points specified in the figure. If banks inject the capital to cover their losses in the scenario, then after doing so, the probability that  $SAD \geq .05$  is  $\leq .05$ . Further details are contained in section 4 of the text.

Figure 8: Can SIR/COP detect the factors out of sample



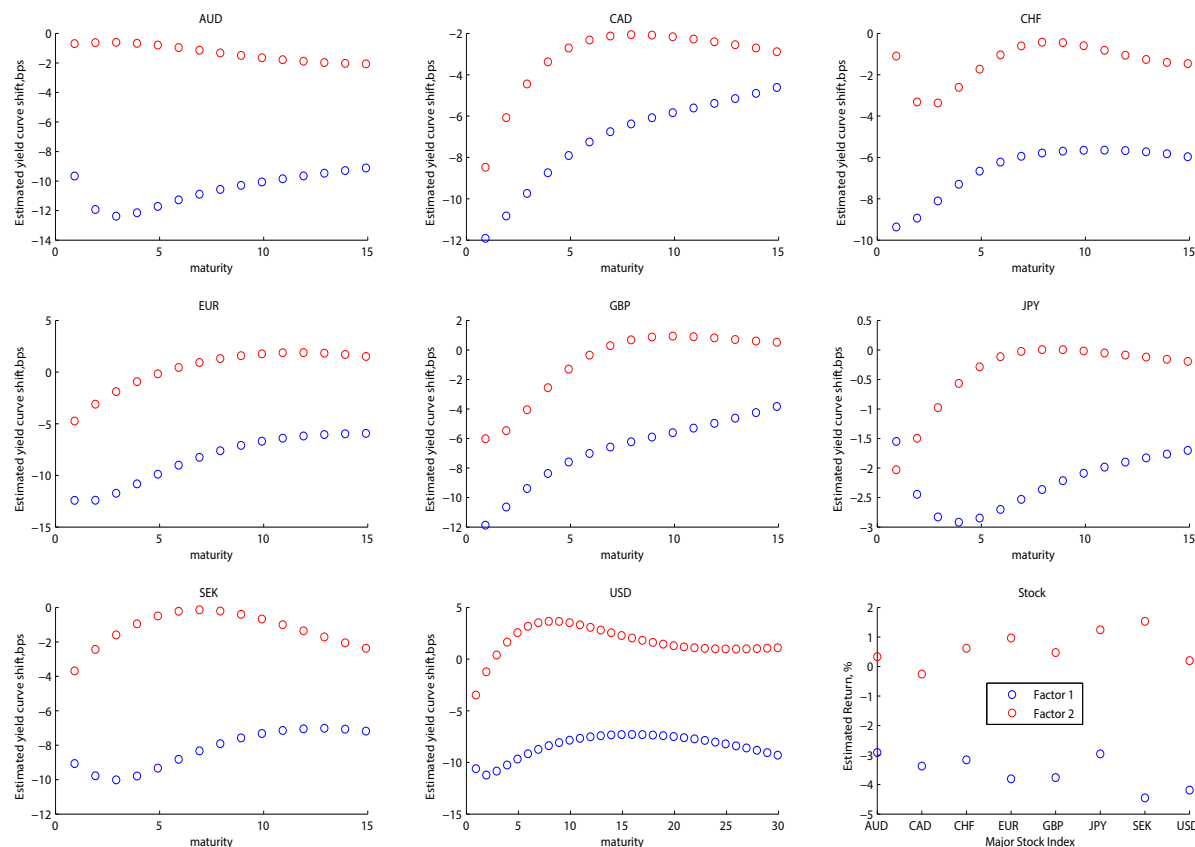
*Notes:* For banks whose portfolios are highly correlated with one of the first 3 principal components of simulated yield curve changes, the figure examines if the factor that COP/SIR identifies using one subsample of the data with 5,000 observations (True F) is highly correlated with the COP/SIR factor estimated using only 500 observations from the same data generating process (F-COP). The fact that the true and estimated factors are highly correlated is evidence that SIR and COP are successfully identifying the true factors that drive systemic risk. For further details, see section 4.2 of the text.

Figure 9: The importance of accounting for portfolio composition when choosing stress-scenarios: A. The effect of one-standard deviation factor shock for a random bond portfolio



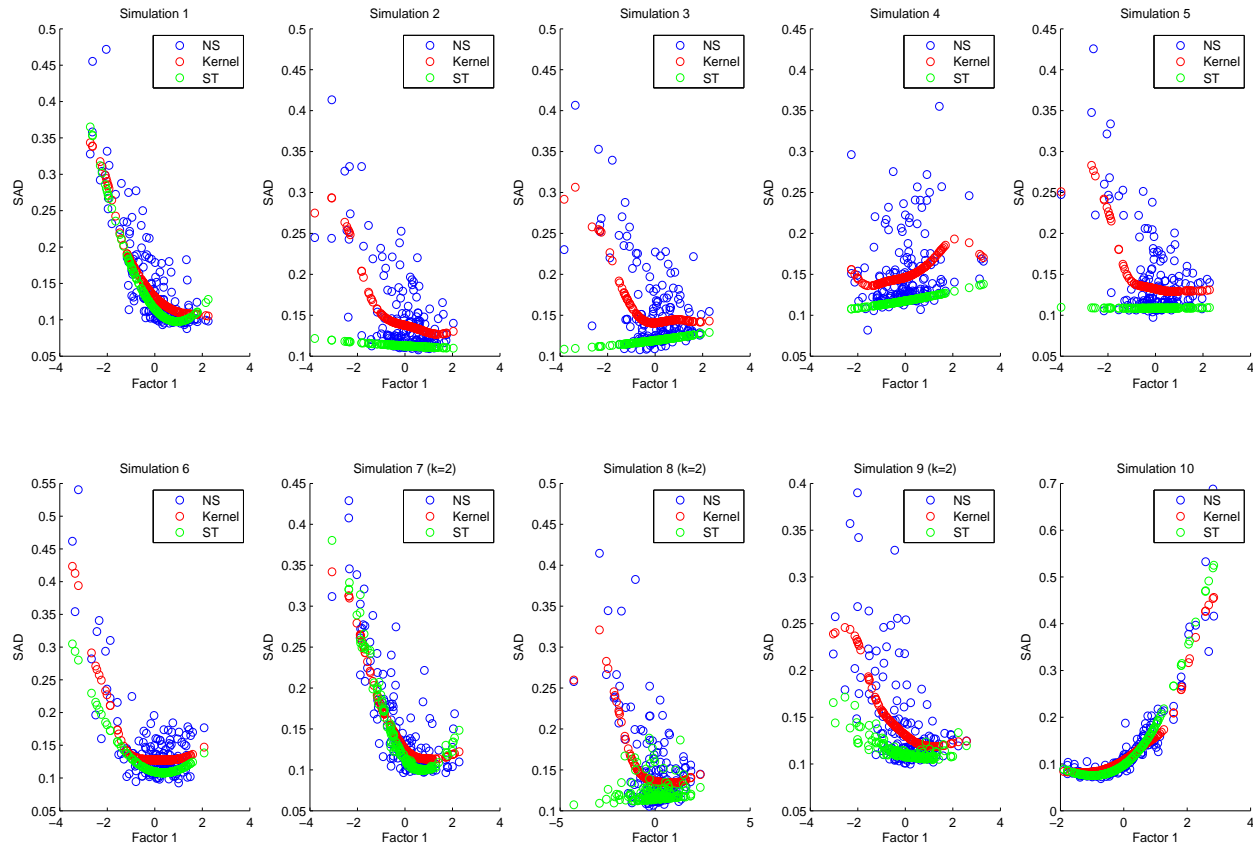
Notes: When 6 banks hold the same portfolio, composed exclusively of bonds, the figure shows the yield curve changes associated with a stress scenario that is created by shocking the single factor for systemic risk that is identified by SIR/COP by one standard deviation. This figure is contrasted with Figure 10. For further details see section 4.2 of the text.

Figure 10: The importance of accounting for portfolio composition when choosing stress-scenarios: B. The effect of one-standard deviation factor shock for a random portfolio of bonds and stocks



*Notes:* When 6 banks hold the same portfolio, that is composed of 50% of the bond portfolio used in Figure 9 and 50% of a random portfolio of country stock indices, the figure shows the yield curve changes and stock return changes associated with a stress scenario that is created by shocking each of the two factors for systemic risk that are identified by SIR/COP by one standard deviation. The most important factor (factor 1) depresses stock prices and steepens yield curves; the second factor (factor 2) mostly changes yield curve curvature. The contrast between the optimally chosen systemic stress scenarios in Figures 9 and 10 show that it is very important to account for portfolio composition when choosing stress-scenarios.

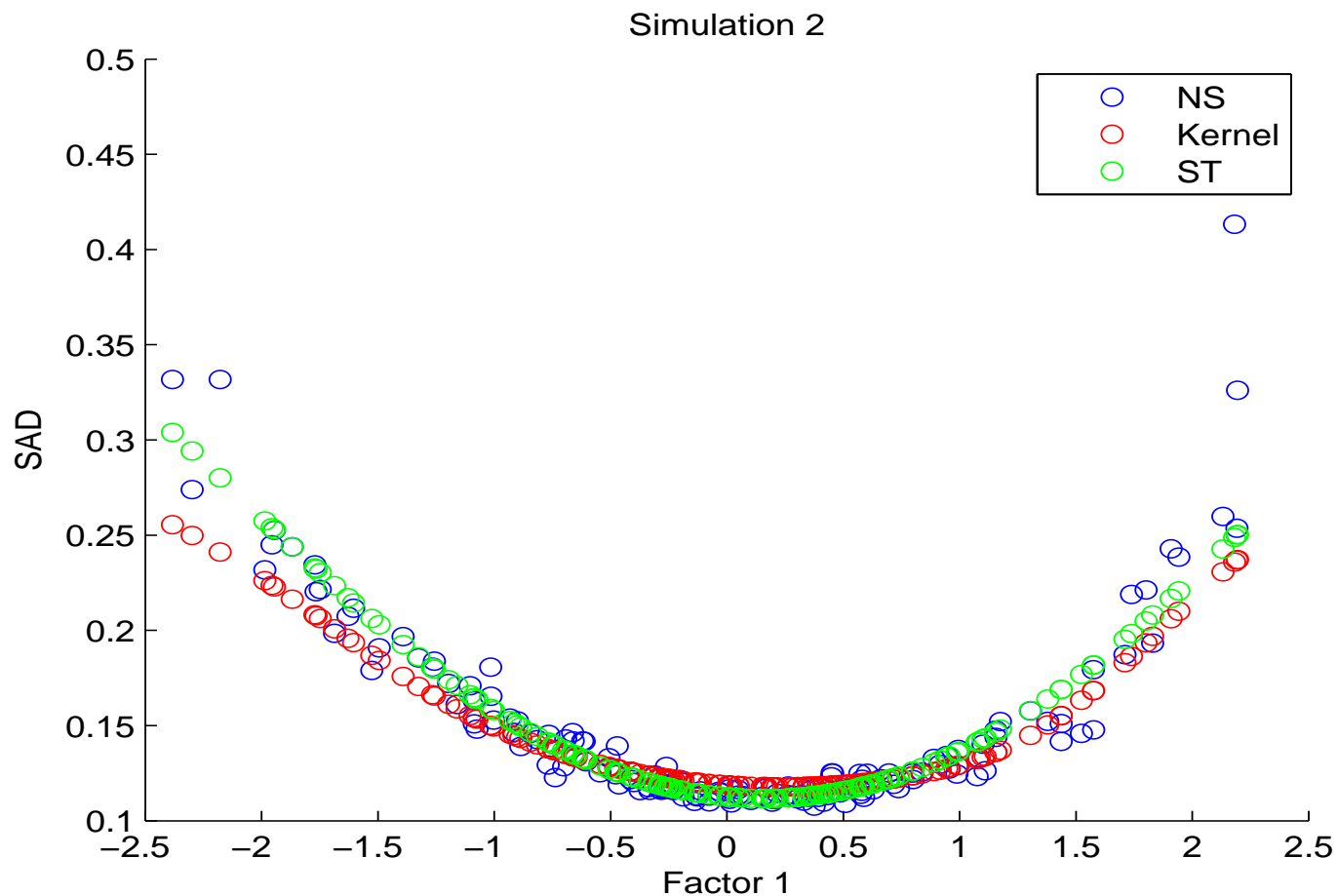
Figure 11: SIR Simulations with occasional symmetry



*Notes:* The figure studies factor identification and stress-scenario creation when banks hold portfolios for which Sliced Inverse Regression is expected to have difficulty identifying systemic risk factors. As noted in the text, SIR can fail to detect factors when SAD is a symmetric function of the  $X$  variables that determine SAD. To study the effects of symmetry, in each panel (labeled simulation 1 - 10) a risky portfolios is randomly created, and each of 6 banks is long or short the same risky portfolio with equal probability. In this setting, SAD can be highly symmetric especially if 3 banks are long and 3 are short the risky portfolio (as in simulation 2). The figure presents the same type of information as Figure 4, but shows that when symmetry is present the extracted factors and stress scenarios created based on those factors do not track SAD well.

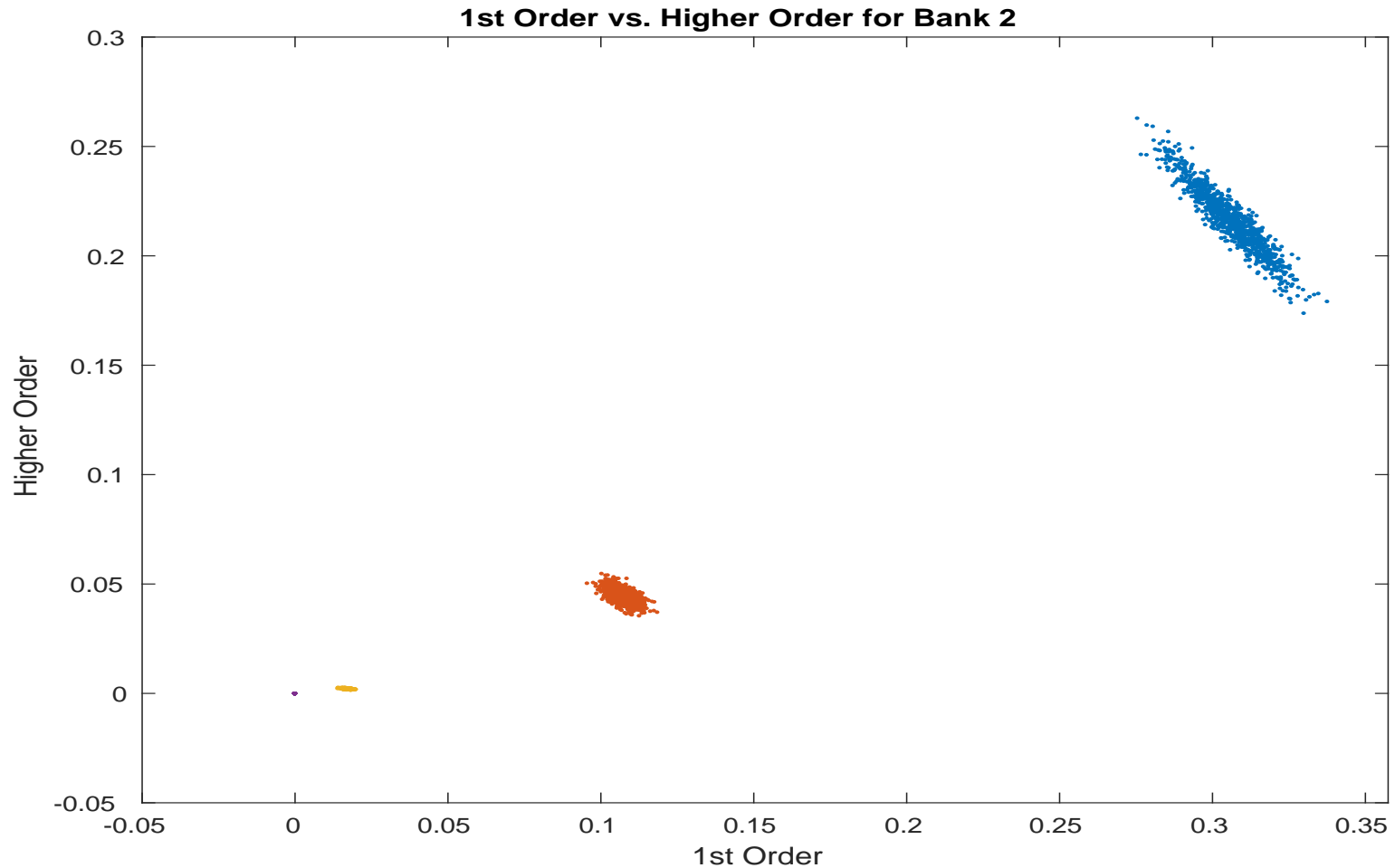


Figure 12: SIR Simulation with symmetry: Scenario 2 – Corrected



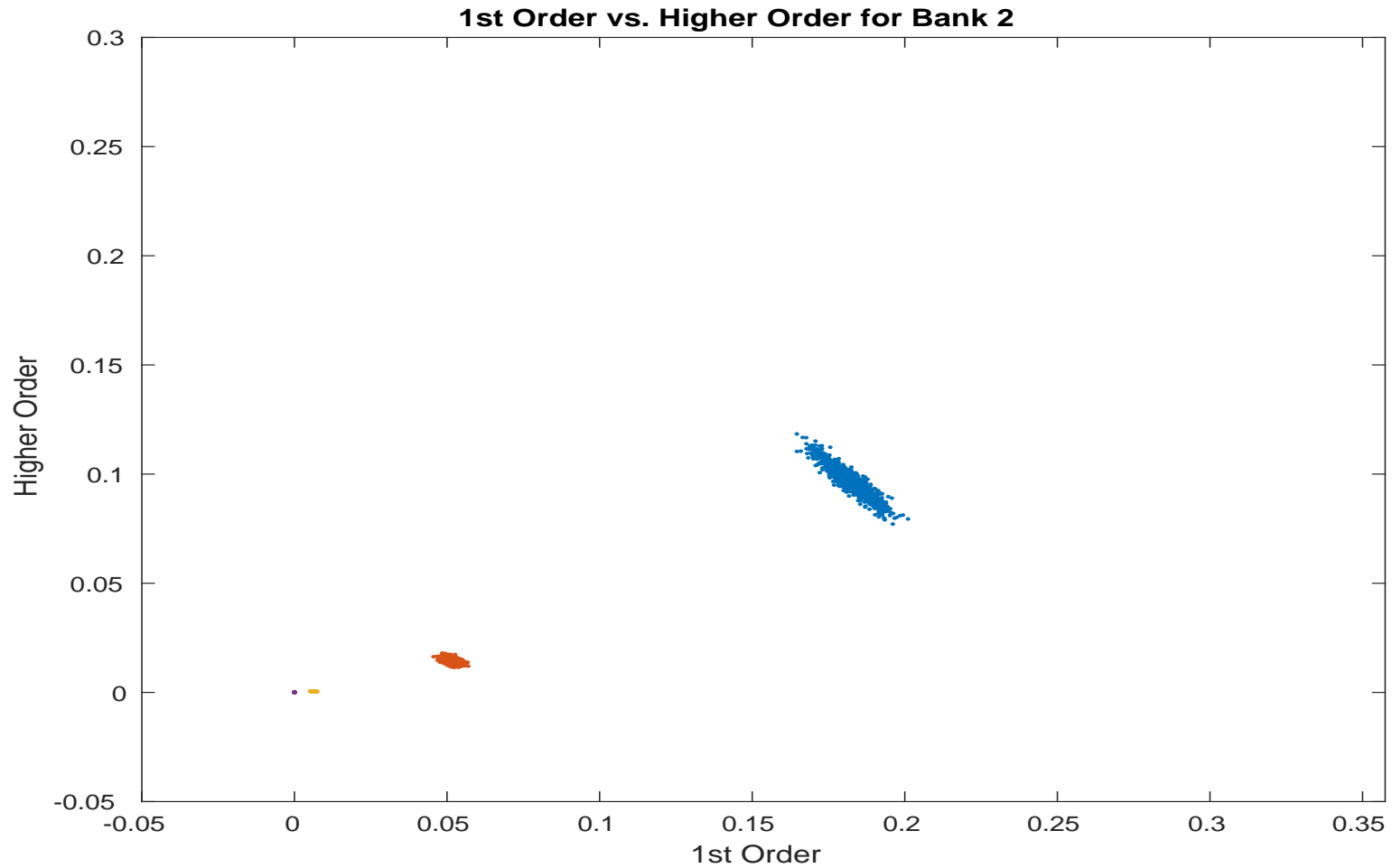
*Notes:* The figure studies factor identification and stress-scenario creation when banks hold portfolios for which Sliced Inverse Regression is expected to have difficulty identifying systemic risk factors. This figure focuses on Scenario 2 from Figure 11 because that scenario does particularly poorly in detecting the sources of systemic risk. To handle those situations a correction is used that extracts systemic risk factors while restricting the factor extraction to banks whose distress functions (a building block of SAD) are positively correlated. This procedure successfully extracts the factors. The figure shows this corrected procedure does a good job of uncovering the true symmetric relationship between the factors and SAD in this circumstance.

Figure 13: Illustration of Network Uncertainty Effects on Bank PD: Baseline Case



*Notes:* The figure is based on Pritsker, 2016, “Network Uncertainty and Interbank Markets.” For a set of 1000 different interbank deposit networks that are drawn from the conditional distribution of the network exposures given public information, the figure illustrates the effect that network uncertainty has on banks probability of default for high (yellow), medium (red), and low returns (blue) on banks non-interbank portfolios. The figure shows that when returns are high, the first order and higher order effects of the networks on the default probability of bank 2 are tightly distributed. However, when returns of banks in the system deteriorate, then the effects of the networks on banks PDs are more spread out (the red and blue points) and consequently uncertainty about the effect of network connections is more important for banks’ risk.

Figure 14: Illustration of Network Uncertainty Effects on Bank PD Following A Capital Injection



*Notes:* The figure is based on Pritsker, 2016, “Network Uncertainty and Interbank Markets.” The figure illustrates the same information as figure 13 except that additional capital has been injected into banks at date 0. The figure shows that when bank 2 and the other banks in the system are better capitalized, then it shrinks the effects of network uncertainty relative to the baseline case in which there was not a capital injection into the banks.